

Causal Interpretations of Probability¹

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The prospects of a causal interpretation of probability are examined. Various accounts both from the history of scientific method and from recent developments in the tradition of the method of arbitrary functions, in particular by Strevens, Rosenthal, and Abrams, are briefly introduced and assessed. I then present a specific account of causal probability with the following features: (i) First, the link between causal probability and a particular account of induction and causation is established, namely eliminative induction and the related difference-making account of causation in the tradition of Bacon, Herschel, and Mill. (ii) Second, it is shown how a causal approach is useful beyond applications of the method of arbitrary functions and is able to deal with various shades of both ontic and epistemic probabilities. Furthermore, I clarify the notion of causal symmetry as a central element of an objective version of the principle of indifference and relate probabilistic independence to causal irrelevance.

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1. Introduction

Research on probabilistic causality has been a thriving enterprise since about the 1980s addressing the mainly methodological question how causality can be inferred from statistical data. By contrast, this article is about causal probability, i.e. the conceptual question how probability can be integrated into a general framework of induction and causation.

In recent discussions on the foundations of probability, a novel class of objective interpretations has been proposed that is distinct from the more familiar propensity and frequency accounts (Strevens 2006, 2011; Rosenthal 2010, 2012; Abrams 2012). The interpretations essentially stand in the tradition of an approach by 19th-century methodologist Johannes von Kries and of related work on the method of arbitrary functions. For reasons that will soon become clear, I subsume these and related approaches under the notion of causal probability. Two common features are particularly important: (i) First, causal interpretations replace or supplement the principle of insufficient reason by an objective version of the principle of indifference³ that refers to physical or causal symmetries. This distinguishes causal interpretations both from frequentist approaches, which exclusively refer to relative frequencies as fundamental evidence for probabilities, and from logical accounts, which base probabilities on ignorance via the principle of insufficient reason, i.e. a purely epistemic reading of the principle of indifference. As we will see, the objective variant of the principle of indifference is not troubled by the central objections brought forth against the principle of insufficient reason, in particular the ambiguities in its application called Bertrand's paradox. (ii) Second, causal interpretations employ a notion of probability in terms of the ratio between favorable *conditions* and all conditions. This is another subtle but crucial difference to frequency interpretations which define probability in terms of the ratio between the number of *events* leading to a certain outcome and the total number of events. As will be shown in Section 3, rendering probability relative to the conditions determining an ensemble or collective provides for a simple solution to a specific version of the reference class problem.

Note that propensity interpretations also frame probability in terms of circumstances or conditions and they sometimes make the link to causation, but the causal approach presented

³ A note on terminology: In the following the term 'principle of indifference' will be used to refer to both an epistemic version, called 'principle of insufficient reason', and an objective version, called 'principle of causal symmetry'.

here differs in important respects. First of all, propensity accounts rely on a distinct ontological category, namely propensities in the sense of tendencies or dispositions. The relation with causality is not always clarified, but if it is, as in Karl Popper's later work, then propensities are mostly considered to be more general than causation. By contrast, the causal interpretation presented in Sections 3 to 6 tries to situate probability within a framework of causal reasoning. While propensity accounts focus conceptually on dispositions or tendencies and rather casually remark upon the parallel with causation, the interpretation proposed here starts with a detailed and specific account of causation and then examines how probability fits into the picture. Furthermore, a number of concepts are central to the causal approach that are not usually evoked in the exposition of propensity interpretations, in particular the notion of causal symmetry leading to an objective version of the principle of indifference (Section 4) and the causal construal of probabilistic independence based on judgments of causal irrelevance (Section 5).

In Section 2, I discuss various proponents of a causal approach to probability from the 19th century as well as more recent developments in the tradition of the method of arbitrary functions. The latter are mainly due to Michael Strevens, Jacob Rosenthal, and Marshall Abrams, and are henceforth abbreviated as SRA-approach. I briefly indicate how causal probability resolves several objections against other interpretations of probability, e.g. the problem of distinguishing between accidental and necessary relations in the frequentist approach, or problems regarding the principle of indifference in the logical approach. I then point out some shortcomings of the SRA-approach. Besides some technical difficulties, it makes no connection with a general framework of induction and causation. Also, it cannot handle indeterminism and epistemic probabilities. Later in the article, I suggest how causal probability can deal with these issues.

Starting from Section 3, I will develop a specific account of causal probability. First, two fundamental inductive frameworks are outlined, enumerative and eliminative induction. For each, I show how probability can be integrated. Enumerative induction leads to a naïve-frequency view of probability, which suffers from the usual problems, in particular that it cannot distinguish between law-like and accidental frequencies. Eliminative induction resolves this issue by carefully keeping track of all circumstances or conditions under which a phenomenon happens. The corresponding account of probability, which distinguishes different types of conditions, is termed causal probability. What I will call the *collective conditions* determine the possibility space of a probabilistic phenomenon, i.e. all possible outcomes. The outcomes are categorized and the classes are labeled, where the labels are called *attributes*⁴. The *range conditions* (together with the collective conditions) then determine exactly which of the attributes occurs, at least in deterministic contexts. While the collective conditions remain constant for a probabilistic phenomenon, the range conditions will vary. A *measure* over the input space, spanned by the range conditions, denotes the limiting relative frequency with which the different input states are instantiated. In principle, this measure is also fixed by the collective conditions. *Causal probability then is calculated as the fraction of input states, weighted with the measure, that lead to a certain attribute.*

⁴ The terms 'attribute' (translated from the German 'Merkmal') and 'range' (German 'Spielraum') are used in reverence to von Mises and von Kries, respectively, on whose ideas the present essay draws substantially.

Rendering probability relative to collective conditions and measure resolves the mentioned technical problems of the SRA-approach while introducing an irreducible epistemic element.

Section 4 introduces the notion of a causal symmetry which allows inferring probabilities without taking recourse to relative frequencies of input states or of outcome events. A causal symmetry basically consists in a relabeling of the outcome space that does not affect the probability distribution. The concept leads to an objective version of the principle of indifference, which I term *principle of causal symmetry*. In the simplest case, two attributes that exhibit a causal symmetry are assigned equal probability. Furthermore, I argue that applications of the epistemic principle of insufficient reason can be reduced to the principle of causal symmetry, whenever the resulting probabilities are predictive. If the relevant causal symmetries are not epistemically accessible, as is often the case, relative frequencies can be consulted as a weaker type of evidence for predictive probabilities.

In Section 5, the notion of probabilistic independence is explicated at some length establishing its relationship with causal irrelevance as determined by eliminative induction. Independence guarantees randomness in the sequence of input states and consequently of attributes. Since many theorems in probability theory like the law of large numbers presuppose independence of trials, a causal construal of independence is another crucial ingredient of the causal interpretation of probability. It broadly corresponds to the notion of randomness in the frequentism and exchangeability in the subjectivist approach. The definition of probability in Section 3b, the principle of causal symmetry, and the causal rendering of probabilistic independence should be considered as a coherent package of the account of causal probability proposed in this essay.

Finally, various ontic and epistemic aspects in probability statements are identified in Section 6, and it is shown how the framework of causal probability can cover a wide range of applications from indeterministic phenomena to probabilities from causal symptoms to the probabilities of hypotheses.

2. Predecessors and contemporary debate

2a. Historical proponents: Cournot, Mill, von Kries

The two main ingredients of a causal interpretation as sketched in the introduction and elaborated later on in the article can be found with a variety of writers until the end of the 19th century. As already indicated, the viewpoint is rather rare in the 20th century presumably due to a widespread hostility towards inductive or causal approaches in science.

The distinction between an epistemic principle of insufficient reason and an objective principle of causal symmetry may be foreshadowed already in Laplace's classic 'Philosophical Essay on Probabilities': "The theory of chance consists in reducing all events of the same kind to a certain number of cases equally possible, that is to say, to such as we may be equally undecided about in regard to their existence, and in determining the number of cases favorable to the event whose probability is sought." (Laplace 1902, 6-7; see also Strevens, Ch. 3.2) Of course, Laplace has in mind what was later called the classical

definition of probability, i.e. the ratio of favorable to all possible cases. But everything hinges on the exact interpretation of equal possibility and how it is determined. Curiously, Laplace alludes both to epistemic and objective aspects, though these are not clearly held apart in his writing. In the quote given above, equal undecidedness implies an epistemic reading of equal possibility. But a later discussion of a loaded die evokes objective connotations in that Laplace distinguishes between judgments with respect to the knowledge of the observer and the presumably objective bias manifest in the coin. Laplace adds that the determination of respective possibilities is “one of the most delicate points of the theory of chances” (p. 11).

Other authors have been more explicit in drawing the distinction between epistemic and objective versions of the principle of indifference. One of the clearest expositions is due to Antoine-Augustin Cournot, who in the following quote delineates a principle of insufficient reason, which cannot establish objective probabilities: “If, in an imperfect state of our knowledge, we have no reason to believe that one combination is realized rather than another, even though in reality these combinations are events that may have unequal mathematical [i.e. objective] probabilities or possibilities, and if we understand by the probability of an event the ratio of the number of combinations that are favorable to the event to the total number of combinations that we put on the same line, this probability could still serve, in lack of a better option, to fix the conditions of a bet [...]; but this probability would not anymore express the ratio that really and objectively exists between things; it would take on a purely subjective character and could vary from one individual to the other depending on the extent of her knowledge.”⁵ (1843, 438, my translation)

Cournot also sketches the role of frequencies with respect to objective probabilities leading to the following colloquial statement of the law of large numbers: “If one considers a large number of trials of the same chance process, the ratio of the number of trials where the same event happens to the total number, becomes perceptibly equal to the ratio of the number of chances favorable to the event to the total number of chances, or what one calls the *mathematical probability* of an event.”⁶ (437, my translation) According to Cournot, the chances are measured in terms of the possibilities that certain conditions occur together to produce a particular type of event. Obviously, he employs a notion of probability distinct from relative frequencies referring to the ratio of favorable to all conditions or circumstances.

Thus, Cournot’s account shows both ingredients of causal probability that were identified in the introduction: the distinction between an epistemic and an objective version of the principle

⁵ “Si, dans l’état d’imperfection de nos connaissances, nous n’avons aucune raison de supposer qu’une combinaison arrive plutôt qu’une autre, quoiqu’en réalité ces combinaisons soient autant d’événements qui peuvent avoir des probabilités mathématiques ou des possibilités inégales, et si nous entendons par probabilité d’un événement le rapport entre le nombre des combinaisons qui lui sont favorables, et le nombre total des combinaisons mises par nous sur la même ligne, cette probabilité pourra encore servir, faute de mieux, à fixer les conditions d’un pari, d’un marché aléatoire quelconque; mais elle cessera d’exprimer un rapport subsistant réellement et objectivement entre les choses; elle prendra un caractère purement subjectif, et sera susceptible de varier d’un individu à un autre, selon la mesure de ses connaissances.”

⁶ “Lorsque l’on considère un grand nombre d’épreuves du même hasard, le rapport entre le nombre des cas où le même événement s’est produit, et le nombre total des épreuves, devient sensiblement égal au rapport entre le nombre des chances favorables à l’événement et le nombre total des chances, ou à ce qu’on nomme probabilité mathématique de l’événement.” (437)

of indifference and a definition of probability that refers to the number of favorable conditions, not instances.

The basic idea of an objective causal interpretation distinct from a frequentist approach is present with several other authors in the 19th century, for example in the writings of John Stuart Mill: “The probability of events as calculated from their mere frequency in past experience affords a less secure basis for practical guidance than their probability as deduced from an equally accurate knowledge of the frequency of occurrence of their causes.” (1886, 355) Mill also recognizes the distinction between an epistemic and an objective reading of the principle of indifference. For example, he criticizes the alleged purely epistemic reading by Laplace: “To be able [...] to pronounce two events equally probable, it is not enough that we should know that one or the other must happen, and should have no grounds for conjecturing which. Experience must have shown that the two are of equally frequent occurrence.” (351) Mill sketches several options how the latter could happen, e.g. for the case of a coin toss: “We may know [that two events are of equal occurrence] if we please by actual experiment; or by the daily experience which life affords of events of the same general character; or deductively, from the effect of mechanical laws on a symmetrical body acted upon by forces varying indefinitely in quantity and direction.” (351) Here, Mill introduces the important distinction between evidence in terms of frequencies and in terms of causal symmetries to establish objective equipossibility (cf. Section 4d). On this basis, he roughly formulates the notion of causal probability referring not to the frequency of events, but to causal conditions: “We can make a step beyond [the frequentist estimation of probabilities] when we can ascend to the causes on which the occurrence of A or its non-occurrence will depend, and form an estimate of the comparative frequency of the causes favourable and of those unfavourable to the occurrence.” (355)

Curiously, Mill eventually retreats from this position that he so clearly formulated in the first edition of his ‘Logic’, adding the following comment in later editions: “I have since become convinced that the theory of chances, as conceived by Laplace and by mathematicians generally, has not the fundamental fallacy which I had ascribed to it [essentially referring to the epistemic reading of the principle of indifference].” (351) Mill claims that probability is fundamentally an epistemic notion and that probabilistic statements have no objective meaning anyways, because in a deterministic world any future event is fully determined by preceding conditions.

It remains somewhat unclear where Mill is heading with these remarks. Does he just want to rehabilitate the epistemic reading of the principle of indifference or does he want to deny the distinction between epistemic and objective readings altogether? From the viewpoint of this essay, Mill is correct that in a deterministic world, there is an epistemic element to any probabilistic statement, but he apparently fails to recognize that a fairly objective meaning of probability nevertheless remains feasible: if one always relates probability *to a causally determined collective* (as elaborated in Section 3b). In any case, it is quite remarkable to observe how even an ingenious thinker like Mill struggles with the concept of probability.

Finally, the approach of Johannes von Kries should be mentioned (as summarized in his 1886, vii-viii). His account was the most influential on 20th-century philosophy, both on discussions

within the Vienna Circle (Waismann, Wittgenstein) and on recent proposals regarding a novel class of objective probabilities (Strevens, Rosenthal, Abrams). Central to von Kries' notion of probability is the *spielraum*⁷ concept denoting the range of initial conditions that lead to a certain result. In principle, probability is determined by the ratio of the measure of the *spielraum* leading to a specific outcome to the measure of the entire *spielraum*. Based on this idea, von Kries formulates three conditions for numerical probabilities: (i) the different possibilities must correspond to *comparable* ('vergleichbar') *spielräume*⁸. In particular, it should be feasible to establish the equality in terms of measure of the various *spielräume* leading to different outcomes. (ii) Furthermore, the *spielräume* should be *original* ('ursprünglich'), i.e. the equality of the *spielräume* must not cease to be the decisive criterion for our expectations when tracing the further history of the conditions making up the *spielräume*. (iii) Third, von Kries requires that the *spielräume* be *indifferent* ('indifferent'), i.e. only the size of the *spielräume* and no other logical conditions should be relevant for the probability. According to von Kries, the most important criterion in this respect is that a small change in conditions already leads to a different outcome. The various outcomes are supposed to alternate rapidly when continuously changing the conditions.

It is mainly this last criterion that establishes the parallel with the *method of arbitrary functions*, a term coined by Henri Poincaré (1912, p. 148). The French mathematician is usually seen as the originator of this tradition, although many ideas are already present in the mentioned work by von Kries (1886; later proponents are von Smoluchowski 1918, Hopf 1936; for a philosophical-historical overview, see von Plato 1983). In general, proponents of the method of arbitrary functions aim to establish objective probability for deterministic phenomena. Building on physical instability, they argue that any sufficiently regular distribution over the initial conditions leads to roughly the same ratio of occurrences in macroscopic outcomes. Primary applications are games of chance like roulette, which already Poincaré discussed in much detail, or the throwing of dice and coins.

Von Kries' account can broadly be classified as causal probability because the two criteria outlined in the introduction are present in his theory as well. First, his treatise on probability contains one of the most insightful assessments of the principle of insufficient reason in the history of probability (1886, Ch. 2). Second, he defines probability not in terms of frequencies of events but in terms of the ratio between different *spielräume*, i.e. conditions.

The outlined accounts are meant to be exemplary, a deeper look into 19th-century discussions on probability would presumably reveal that similar causal viewpoints were widespread.⁹ In the first half of the 20th century, the ideas of von Kries were picked up and developed into an objective interpretation of probability by Friedrich Waismann (1930/1931), who claims in turn to have been influenced by Wittgenstein.¹⁰ These accounts are somewhat similar to independent suggestions elaborated in recent years by Michael Strevens, Jacob Rosenthal, and Marshall Abrams, to which we will turn now.

⁷ *Spielraum* translates to 'range of possibilities'.

⁸ I am using the German plural *Spielräume*.

⁹ In fact, already Jacob Bernoulli in his *Ars conjectandi* interpreted equipossibility in a causal manner: "All cases are equally possible, that is to say, each can come about as easily as any other" (1713, 219; cited in Hacking 1971, 344).

¹⁰ For a historical overview, see Heidelberger (2001).

2b. Contemporary debate: Abrams, Rosenthal, Strevens

Apparently, the history of causal interpretations of probability before the 20th century is quite rich and it seems plausible that the demise of this perspective more or less parallels the rise of causal skepticism in the beginning of the 20th century. At the same time, the distinction between frequentist evidence for objective probabilities and evidence in terms of causal symmetries largely disappears from the debate leading to a purely frequentist view of objective probabilities. Furthermore, only the epistemic version of the principle of indifference remains as a centerpiece of the logical interpretation, while the objective reading is largely abandoned. A notable exception in the latter respect are the writings of John Maynard Keynes who clearly recognizes a difference between ascribing equal probabilities on the basis of no evidence as opposed to evidence in terms of frequencies or relevant circumstances. He believes that the distinction is gradual and introduces the notion of *weight of argument* to account for it (1921, Ch. VI). But the idea has not caught on in 20th-century literature on probability.¹¹

In recent years, one can observe a revival of objective interpretations that go beyond the frequency account by making explicit reference to initial conditions as well as system dynamics and thus bear resemblance to the historical accounts depicted in the previous section. This type of objective interpretations, which has been more or less independently developed by Marshall Abrams, Jacob Rosenthal, and Michael Strevens, substantially relies on ideas from the method of arbitrary functions.¹²

The best-known account in this modern tradition is Michael Strevens' *microconstant probability* (2011; see also 1998, 2006, 2013). In part, his approach is inspired by Maxwell's derivation of the molecular velocity distribution in an ideal gas which was carried out without empirical data about those velocities, i.e. without frequency data. Strevens elaborates in much detail the distinction between an objective and an epistemic reading of the principle of indifference (2013, Ch. 3). In his recent book 'Tychomancy', he lays out the most important principles of reasoning based on the objective version, which he terms *equidynamics*, by analyzing exemplary processes such as stirring or shaking (Ch. 5-8).

In one recent article, Strevens defines microconstant probability as an objective physical probability for deterministic systems along the lines of the method of arbitrary functions: "The event of a system S's producing an outcome of type e has a microconstant probability equal to p if (1.) the dynamics of S is microconstant with respect to e, and has strike ratio p, (2.) the actual initial conditions of nearly all long series of trials on systems of the same type as S make up macroperiodically distributed sets, and (3.) the macroperiodicity of the initial conditions is robust." (2011, 359)

Apparently, the crucial notions are *microconstancy* and *macroperiodicity*. The former refers to the condition that "within any small but not too small neighborhood, the proportion of

¹¹ As Keynes himself stated, he was influenced by von Kries in framing the notion of weight of argument (cp. Fioretti 1998).

¹² One should also mention the work of Richard Johns, who proposed a causal account of chance: "the chance of an event is the degree to which it is determined by its cause" (2002, 4). Moreover, propensity accounts are related to the causal approach, as already pointed out in the introduction and discussed further in Section 3c.

initial conditions producing a given outcome is [approximately] the same” (2013, 11). This proportion is called *strike ratio* and it essentially determines the probability modulo substantial problems concerning the limiting process to infinitesimal neighborhoods and thus to exact probability values. Macroperiodicity denotes a certain smoothness in the probability distribution over initial conditions, such that neighboring initial conditions leading to different results should occur with approximately the same frequency in long series of trials.¹³ This uniformity together with microconstancy leads to stable strike ratios and thus probabilities that are largely independent of the exact probability distribution over initial conditions. Finally, robustness in Strevens’ third premise refers to counterfactual robustness, i.e. that counterfactual and predictive statements about frequencies are sufficiently reliable. Typical applications for microconstant probability are games of chance like roulette or playing dice, but Strevens believes that the notion also covers scientific applications from statistical physics¹⁴ to the theory of evolution. Obviously, Strevens’ approach features both characteristics of causal probability mentioned in the introduction.

A further prominent account in the tradition of the method of arbitrary functions is Marshall Abrams’ *far-flung frequency (FFF) mechanistic probability* (2012). His approach, although independently developed, bears close resemblance to the accounts of both Strevens and Rosenthal. In particular, he relies on the same concepts of microconstancy and macroperiodicity as coined by Strevens. Abrams introduces the notion of a *causal map device*, which maps the input space to the outcome space, and partitions the outcome space into basic outcomes. A bubble is defined as a region in the input space containing points leading to all possible outcomes. A partition of the entire input space into bubbles he calls a bubble partition. Probability then is determined in the following manner: “There is a bubble partition of the [causal map] device’s input space, such that many ‘far flung’ large natural collections of inputs together determine an input measure which makes most of the collections macroperiodic (and such that moderately significant changes in the spatiotemporal range across which natural collections are defined don’t significantly affect outcome probabilities).” (Sec. 6) For lack of space, I won’t go into details what exactly Abrams understands by “‘far flung’ large natural collections of inputs”, but essentially they fulfill two conditions: they are microconstant and they reflect actual input patterns in the world (Sec. 4.2). Abrams emphasizes that he intends an objective probability interpretation that can account for a wide range of applications in games of chance, statistical mechanics and perhaps also the social and biological sciences (Sec. 6).

Finally, Rosenthal presents a very clear and thoroughly argued account of an objective probability interpretation largely construed around the notion of arbitrary functions, which he terms *natural range conception* in reminiscence of von Kries’ *spielraum*-concept. He formulates two equivalent versions, one in terms of an integral over those initial states that lead to a desired outcome and the other referring to the ratio of ranges in the initial state space. I will focus on the second explication, which Rosenthal frames as follows: “Let E be a random experiment and A a possible outcome of it. Let S be the initial-state space attached to

¹³ In ‘Tychomancy’, Strevens replaces the term by the “in essence identical” (2013, 58) notion of *microequiprobability* that the probability density is approximately uniform over any small contiguous interval or region in the initial state space (2013, 246).

¹⁴ For a related discussion, cf. Myrvold 2011.

E , and S_A be the set of those initial states leading to A . We assume that S and S_A are measurable subsets of the n -dimensional real vector space \mathbf{R}^n (for some n). Let μ be the standard (Lebesgue-)measure. If there is a number p such that for each not-too-small n -dimensional (equilateral) interval I in S , we have

$$\frac{\mu(I \cap S_A)}{\mu(I)} \approx p$$

then there is an objective probability of A upon a trial of E , and its value is p .” (2012, 224)

Thus, Rosenthal explicitly frames his account as an objective probability interpretation for deterministic systems (2010, Sec. 5.3). In summary, the idea is that the probability of an outcome is proportional to that fraction of the initial-state space leading to the outcome, as determined by the Lebesgue measure. Since Rosenthal aims to develop an account for deterministic chance, i.e. he wants to eliminate epistemic aspects as far as possible, he has to require that in the initial-state space the conditions leading to the different outcomes are everywhere equally distributed, at least when looking with sufficient coarse-graining. This implies that any sufficiently smooth density function over the initial-state space will lead to approximately the same probability, which establishes the connection to the approach of arbitrary functions and the close relationship with Strevens’ microconstant probability relying on the notions of microconstancy and macroperiodicity. Of the three accounts discussed in this section, Rosenthal’s definition remains closest to the original ideas of von Kries’ spielraum conception by referring explicitly to a specific measure over the initial space. Otherwise, the method of arbitrary functions could also be understood in terms of a frequentist approach with respect to the occurrence of initial states.

Rosenthal discusses a central objection against his own approach which comes in two slightly differing versions (2012, Sec. 4; 2010, Sec. 5.5). First, an eccentric distribution over initial states might be realized in nature leading to observed frequencies deviating substantially from p . Rosenthal suggests that at least in some such cases a nomological factor has been overlooked that determines the eccentric distribution. According to the second variant of the objection, there usually exist various ways in which the initial-state space could be reformulated such that it loses the characteristics required for Rosenthal’s definition of probability. In particular, the Lebesgue measure might cease to be an appropriate choice to account for observed frequencies. Thus, one has to motivate why a certain formulation of initial conditions suitable for the natural-range conception is superior to others that are not suitable. Rosenthal essentially acknowledges that these are open problems for his approach.

Note that they are equally troublesome for Strevens’ and Abrams’ account since the concepts of microconstancy and macroperiodicity already presuppose a choice of measure. As a solution, Strevens suggests to always use standard variables, measured in standard ways. Because these tend to be macroperiodically distributed, microconstancy with respect to standard variables is meaningful. While Strevens’ account is quite sophisticated in this respect (2006, Sec. 2.5; 2013, Ch. 12), I believe that the rejoinder eventually fails due to the blurriness and context-dependence of the notion of standard variable. After all, most

phenomena can be accounted for in a large number of ways and it is just not plausible that all formulations will always yield microconstancy and macroperiodicity to the same extent.

A related problem concerns the various imprecisions and approximations figuring in the definition of probability of all three accounts. For example, Rosenthal's definition refers to "not-too-small" intervals and the ratio of ranges only approximately determines the probability " $\approx p$ ". In fact, the strike ratio will in general slightly fluctuate between different regions of the initial-state space. Thus, all theorems concerning microconstancy and macroperiodicity also hold only approximately. Especially, when aiming at a purely objective interpretation, these features are troublesome.¹⁵ In Section 3b, I suggest how the outlined technical problems can be avoided by rendering probability measure-dependent.

Due to the close similarity of the accounts developed by Strevens, Rosenthal, and Abrams, I will in the following refer to them as the SRA-approach to objective probability.

2c. Some praise

The causal approach referring to initial or boundary conditions can resolve a number of problems for traditional accounts of probability. These issues are discussed extensively by the authors mentioned in the previous section, so I will not delve into details. Let me just briefly comment on a few points.

According to Strevens, the "fundamental flaw" of the frequency account is that it cannot distinguish between meaningful and arbitrary frequencies and thus cannot reliably ground counterfactual statements and predictions in probabilities (2011, Sec. 2). The issue largely parallels the standard problem of induction. In reply, the causal account offers as a criterion that frequencies are only meaningful, when they result from a collective determined by causal conditions. Of course, this solution can only get off the ground given a defensible notion of causation, a topic that will be addressed in Section 3.

A further major advantage in comparison with frequency theories is that causal interpretations can establish probability independently of observed frequencies, for example by referring to symmetries or by rendering probabilistic phenomena largely independent of the probability distribution over initial states. Among other things, this allows for a non-circular reading of the law of large numbers (e.g. Abrams 2012, Sec. 1.1; Rosenthal 2010, Sec. 5.2).

By relying on some version of the principle of indifference, causal probabilities bear resemblance to logical interpretations of probability. However, the principle of insufficient reason referring to ignorance, as it is used in the logical approach, is notoriously flawed by challenging objections—in particular Bertrand's paradox, which highlights ambiguities in the application of this principle (Bertrand 1889; van Fraassen 1990, Ch. 12). The causal approach resolves these ambiguities by introducing an objective variant of the principle of indifference, later referred to as principle of causal symmetry in the specific account of causal probability to be developed in Sections 3 to 6 (cp. esp. Section 4c).

¹⁵ In personal communication, Michael Strevens has suggested as a response to consider microconstant probability as objective, but slightly indeterminate.

2d. Critical reflections

While clearly being a major step in the right direction, the recent attempts to develop an objective account of probability in the tradition of the method of arbitrary functions suffer from a number of shortcomings. First, there are the technical objections already pointed out towards the end of Section 2b. In addition, there are two more general issues, which I will delineate in the following.

First, the objective accounts of the SRA-approach mostly fail to clarify the relation to epistemic probabilities and therefore implicitly subscribe to an in my view misguided sharp distinction between ontic and epistemic probabilities. Instead, I will pin down in Section 6 several epistemic features that can but need not be present in the assignment of probabilities. I will sketch how the various shades of epistemic and ontic probabilities can all be accounted for in terms of a single causal interpretation. Thus, the range of application is widely extended beyond cases in which the method of arbitrary functions can be employed.

Second, the mentioned accounts all rely on physical or causal laws determining the dynamics of the considered phenomena largely without explaining the origin of these laws. In the worst case, they need to be established inductively leading us back to the problem of distinguishing between meaningful and arbitrary relations, which the SRA-approach aimed to resolve in the first place. Thus, a major task for any approach to probability is to clarify how it fits into a more general framework of induction and causation. This will be attempted in the following.

3. Induction, causation, and probability

In the previous section, a shortcoming of the SRA-approach was identified that the probabilities rely on physical knowledge in terms of dynamics and laws of motion but fail to make a connection with a specific account of induction and a corresponding notion of causation. In the following, I try to ameliorate the situation by comparing two distinct accounts of induction, namely enumerative and eliminative, and by examining how in each case a notion of probability could be integrated. Enumerative induction leads to a naïve frequency account of probability that must be rejected in particular for failing to draw a distinction between accidental and lawlike regularities. By contrast, eliminative induction offers a solution to this problem in terms of a difference-making account of causation, while of course some amount of uncertainty remains for any inductive inference. Trying to implement probability in eliminative induction will lead to an account of causal probability that resembles those presented in Sections 2a and 2b. From now on, the terms ‘causal interpretation of probability’ and ‘causal probability’ more narrowly refer to the specific account to be developed in the remainder of the essay.

3a. Enumerative induction and the frequency theory

Enumerative induction is the rather naïve view that general laws can be deduced from the observation of mere regularities: If in all observations, one finds two events, objects, properties, etc. A and B always conjoined then there supposedly exists a causal connection

between A and B. This basic idea is shared by all naïve regularity conceptions of natural laws and causation.

The generalization to statistical laws is straight-forward although some technical complications arise due to possible fluctuations in the observed frequencies. Basically, if in a sequence of events of type A one finds a more or less constant ratio p for another type of event B, then one can conclude to a statistical law connecting A and B with probability p . For example, if a coin lands heads in approximately one half of all trials, then the probability of this event probably is somewhere close to one half. Serious problems arise because the true value of the probability is usually identified with the limiting frequency in an infinite number of trials. The naïve frequency view thus grants epistemic access only to observed frequencies but not to the underlying probabilities themselves. Therefore, it exhibits considerable difficulties dealing with cases, where the frequencies by pure coincidence deviate from the actual probabilities.

However, at this point we can neglect the problems arising in this regard since the naïve frequency view falls prey to a much more fundamental flaw, the same as the naïve regularity conception of laws and causation: it cannot distinguish between accidental and lawlike statistical relationships, i.e. between those that can ground predictions and successful manipulations and those that cannot (cp. Strevens 2011, Sec. 2; as already discussed in Section 2c). For example, the naïve frequency view cannot handle the following situation of an exchanged coin. Consider a sequence of throws, during which the coin is exchanged at some point with another one looking very much alike. Presumably, the naïve frequentist would have to derive predictions about future events from the whole sequence. He cannot make the crucial distinction between the case, where both coins are structurally similar, and the case, where the coins are structurally distinct, e.g. one fair the other loaded. As we will see shortly, such distinctions can be systematically established only within the context of eliminative induction. In other words, the naïve frequency view leads to an essentially unresolvable reference class problem since it lacks clear rules how to determine structural similarity.

In comparison, the causal interpretation elaborated in this essay accepts that any single event can be attributed to different collectives, which in general imply different probabilities for the event. In other words, there is an ambiguity in the choice of reference class, which however is not fatal to the causal interpretation, since causal probability is defined with respect to a collective. This dissolves what Alan Hájek has termed the metaphysical reference class problem (2007). Note that an epistemic agent acting on the basis of probabilities should use the collective that is as specific as possible in terms of causally relevant conditions under the additional constraint that the agent has epistemic access to some evidence for the corresponding probabilities in terms of symmetries or relative frequencies. By contrast, the fatal reference class problem for the naïve frequentist is that he may construct an ensemble of seemingly similar events, which are however structurally dissimilar, and therefore the resulting frequencies are not predictive. This problem is avoided in the causal approach because the collective conditions are by definition causally relevant for the considered phenomenon and must remain constant during all trials, while the range conditions are supposed to vary randomly.

3b. Eliminative induction and the causal conception of probability

Eliminative induction is distinguished from enumerative induction in that it examines not the mere repetition of phenomena but rather phenomena under varying circumstances or conditions. Eliminative methods determine the causal relevance or irrelevance of conditions for a certain phenomenon. The main methods are the method of difference and the strict method of agreement. The first establishes causal relevance of a condition C to a phenomenon P from the observation of two instances which are alike in all conditions that are causally relevant to P except for C. If in one instance both C and P are present and in the other both C and P are absent, then C is causally relevant to P. The strict method of agreement establishes causal irrelevance in much the same manner, except that the change in C has no influence on P.¹⁶ According to this view of eliminative induction, causal (ir-)relevance is a three-place notion: Condition C is causally (ir-)relevant to P with respect to a background B consisting of further conditions that remain constant if causally relevant to P or that are allowed to vary if causally irrelevant. For further details and a brief elaboration of the related difference-making account of causation, see Pietsch (2014).

How does probability fit into this picture of induction? Note first that both principal methods of eliminative induction presuppose determinism, i.e. that P is fully determined by causal conditions (Pietsch 2014, Sec. 3f). Consequently, we will in the following delineate an essentially epistemic probability conception for deterministic phenomena, while indeterministic probabilities can be integrated later on, as discussed in Sections 6a and 6b.

In developing a probability concept for eliminative induction, the focus must lie on the variation of conditions, which concerns the crucial change in perspective with respect to enumerative induction—much in the spirit of Federica Russo’s variational epistemology for causation (e.g. Illari & Russo 2014, Ch. 16). In particular, a careful distinction between various types of circumstances or conditions needs to be introduced.

We are interested in the impact of a number of potentially relevant conditions C1, ..., CM on a statistical phenomenon P with respect to a constant background B. Since P is statistical, it must be linked to a space O of possible outcome states, which may be continuous and many-dimensional, but will for the sake of simplicity from now on be assumed as discrete and one-dimensional. No additional conceptual difficulties arise in the former case. The outcome space is divided into mutually exclusive regions covering the whole space. These regions are labeled and the labels are called *attributes* M1, ..., MN.¹⁷ Note that the labels are introduced in addition to the parameters spanning the outcome space. They do not constitute causally active conditions for reasons that will become clear later on when the notion of causal symmetry is defined.

Let me now introduce various types of conditions, in particular the distinction between *collective*¹⁸ *conditions* and *range*¹⁹ *conditions*. Both types of conditions are causally relevant

¹⁶ Note that as a complication, causal irrelevance can only be established up to a certain degree corresponding to measurement accuracy.

¹⁷ In reverence to von Mises who used the German term ‘merkmal’ that translates to feature, attribute, characteristic.

¹⁸ Again, we rely on the terminology of von Mises.

(in the sense of difference-making) to P . When examining a particular probabilistic phenomenon, the collective conditions must remain constant, while the range conditions are allowed to vary. The collective conditions fix the occurrence of the class P but do not determine which of the attributes M_1, \dots, M_N will actually happen, i.e. these conditions determine the probability space regarding the various manifestations of the phenomenon P . Note that the collective conditions include the background B .

The range conditions determine the input space of possible input states S_1, \dots, S_O . Via a causal mapping $S \xrightarrow{c} O$, every input state determines which event MX of the M_1, \dots, M_N will actually happen. Thus, collective and range conditions together deterministically fix the event MX . Again, we assumed a discrete, one-dimensional input space for the sake of simplicity, a generalization would add no further difficulties.

In order to derive probabilities, a probability measure W needs to be introduced over the input space, i.e. all possible input states. In principle, this measure is determined by the collective conditions as further discussed in Section 4. It is normalized over the whole input space and determines the frequencies with which the input states of a specific probabilistic phenomenon will be instantiated. Sometimes, when it is possible to clearly specify the process determining the measure, it may make sense to distinguish two types of collective conditions: *set-up conditions* determining the range of possible input states and therefore the outcome space; and *measure conditions*, which fix the measure over the input space and thus the probabilities of the outcomes.²⁰ Note that in the exceptional case of indeterministic phenomena, a measure can be dispensed with and the probabilities directly result from the system's dynamics.²¹

This leads to the notions of a *probabilistic phenomenon* and of *causal probability*:

A probabilistic phenomenon P is determined by collective conditions C , range conditions R spanning the input space S , a probability measure W over the input space and a causal mapping $S \xrightarrow{c} O$ of the input space on the outcome space O . The causal probability of a specific attribute MX , combining a set of possible outcomes of the phenomenon P , is given by the fraction of input states weighted²² with the measure W leading to outcome MX .²³

Thus, probability is always relative to collective conditions and a measure over the input space—which is very much in the spirit of von Mises' famous statement “first the collective—then the probability” (1981, 18).²⁴ Sometimes, when input space and measure are not explicitly known, one may express a probabilistic phenomenon in terms of the attribute space, but a constant collective in terms of collective conditions and measure is nevertheless

¹⁹ The terminology here is of course in reverence to von Kries.

²⁰ There is often a normative component to the measure and thus also to the collective conditions, since it is partly a matter of choice which events to include in a collective and which not.

²¹ We will return to this in Section 6a.

²² Henceforth, I will speak of the ‘weighted fraction of input states’.

²³ I claim that this is the notion of probability that many of the classical thinkers mentioned in Section 2a had in mind. Strevens (2006) makes a similar suggestion, but sees it as a special kind of probability, namely ‘complex probability’, in contrast with ‘simple probabilities’ that appear in or depend on fundamental laws of nature.

²⁴ “we shall not speak of probability until a collective has been defined” (ibid.)

always required. Note also that the basic axioms of probability will be satisfied since the definition is based on fractions referring to a normalized measure.

In Sections 4 and 5, I will introduce several further concepts that are central to the causal interpretation. The notion of causal symmetry and the related principle of causal symmetry, as explicated in Section 4, allow establishing the measure to an extent that the probability distribution of the attributes can be fixed without relying on relative frequencies of initial conditions as evidence. In Section 5, a causal construal of the notion of independence will be given ensuring that any sequence of initial states will be random. Without independence, one could hardly speak of a probabilistic phenomenon, since many theorems of probability theory like the law of large numbers justifying the convergence of relative frequencies to the actual probabilities rely on independence of subsequent events. The definition of probability given above, the principle of causal symmetry, and a causal construal of the notion of independence should be seen as one package making up causal probability.

The connection with eliminative induction can also be understood in terms of a coarse-grained formulation. Instead of examining particular instances, where specific S and O are realized, statistical phenomena P as a whole could be considered determined by certain collective conditions and an attribute distribution, e.g. an ideal gas in a box or a long sequence of throws with a die. The causal relation how changes in collective conditions impact on statistical phenomena can be established by the method of difference and the strict method of agreement presupposing determinism. Predictions and counterfactual statements about probabilistic phenomena can thus be derived by ordinary eliminative induction.

Let me illustrate the proposed notion of probability with the simple example of the throw of a coin (P). The attributes partitioning the outcome space are heads-up (M1) or tails-up (M2). The collective conditions are the causal conditions of the set-up, e.g. concerning the type of coin, the allowed types of throwing, the types of surface on which the coin lands, etc. These conditions are held fix in all instances of the phenomenon. The range conditions are also causally relevant to the outcome but randomly vary from throw to throw: including the exact initial state of the coin before the throw, the initial speed, direction, and torque of the throw, etc. Assuming determinism, the attribute is fixed by the range conditions, corresponding to a mapping $S \xrightarrow{c} O$. Finally, the measure W determines the limiting relative frequency, with which the various range conditions occur. In principle, W is also fixed by the collective conditions, i.e. the instructions how to throw the coin should determine how often certain input states occur. It should be added that it generally suffices when the instructions determine the frequency of initial states to an extent that the attribute distribution is fairly stable. In other words, the measure is seldom fixed to full extent. This is the lesson learned from the method of arbitrary functions. Note finally that the range conditions can usually be formulated in different ways for a probabilistic phenomenon, which requires a complementary adjustment of the measure.

As long as the collective for the throws remains the same and the initial states vary sufficiently, long-run frequencies will almost always closely approximate the actual probabilities according to the mathematical theorem called the law of large numbers. This solves the problem of the exchanged coin of Section 3a. As long as both coins are structurally

similar, e.g. fair, the collective conditions stay the same when the coin is exchanged, and therefore predictions based on combined frequencies can be expected to hold. If one coin is fair and the other loaded, then the instances do not form a collective, because a causally relevant condition has changed and therefore predictions based on relative frequencies will in general fail to hold (though there may be ways of formulating a combined collective, see Section 6b).

Another classic application of probability concerns population statistics, e.g. the question whether a certain person will die at a given age. Regarding this type of problem Mill has claimed that probability lacks an objective meaning since for every individual death is supposedly a matter of deterministic fact (cf. Section 2a). With respect to single-case probabilities in deterministic settings, this assessment is certainly correct. However, there is a fairly objective meaning to probability if relating it to a specific collective and a measure over input states as required by the definition of causal probability given above (regarding a discussion of various epistemic elements in causal probabilities, cf. Section 6).

To determine the probability whether someone will die at a specific age we thus first have to fix a collective specifying causally relevant circumstances, for example the gender of a person, certain habits, e.g. whether he/she smokes, is active in sports, or has pre-existing diseases. The collective conditions leave open the two possibilities of interest that the person dies at a given age or not. Probabilities result from the range conditions, a measure over the input space spanned by the range conditions, and a causal mapping of the input space on the outcome space, although these need not—and often cannot—be made explicit. While admittedly it is impossible to list all the relevant causal conditions for phenomena with a complex causal structure like the death of a person, in principle the construction of a collective according to the definition above is possible assuming determinism. And the fact that insurance companies manage to arrive at fairly stable probability distributions suggests that they have some epistemic access to appropriate collectives.

In combination, collective and range conditions determine whether a person will die or not. Of course, the exact boundary between collective and range conditions is usually quite arbitrary. In the case of population statistics, the collective is mostly determined by choosing a certain group of the total population, for example white male living in New York State. Since epistemic access to causal symmetries is implausible for phenomena of such complexity, the required information about range conditions and measure is derived from past frequency data—under the assumption that this data is representative of the group and that collective conditions and measure will approximately stay the same for the time period that is to be predicted. Note again that the collective should generally be chosen in such a way that it includes all conditions that are known to be causally relevant in a considered instance, if one wants to act on the basis of the resulting probabilities. For example when someone is known to have prostate cancer, this information should be included in the collective conditions concerning an imminent death, if, of course, there is also sufficient frequency data to determine the corresponding probabilities.

3c. A brief comparison with other accounts

In the introduction, I had already pointed out the main differences between the causal approach and the logical as well as the frequentist accounts. With respect to the former, the causal approach relies on an ontic and not on an epistemic version of the principle of indifference. With respect to the latter, the causal approach defines probability in terms of the ratio of favorable boundary or initial conditions and not in terms of relative frequencies of events.

The account proposed in Section 3b is conceptually closest to the SRA-approach and to the propensity theory. It is therefore worthwhile to briefly address the most important differences in each case. With respect to the SRA-approach based on the method of arbitrary functions, a crucial difference is that causal probability²⁵ is always relative to the measure over the input space while the SRA-approach tries to establish that probabilities are independent of the choice of measure. Rendering probability relative to the measure resolves in a simple manner the central objection against the natural-range conception that was described towards the end of Section 2b. Concerning the first situation, i.e. the problem of eccentric distributions over initial states, the causal perspective is the following. If the collective conditions determine an eccentric distribution, the measure must reflect this distribution. By contrast, if an eccentric sequence of initial conditions occurs by coincidence given a non-eccentric measure, then the eccentric sequence must be attributed to chance.

The second situation, Rosenthal worries about, is that reformulations of the initial conditions lead to a change in probabilities. Indeed according to his natural range conception, which relies on the Lebesgue measure over the initial-state space, reformulations could easily imply probabilities in contradiction with observed frequencies. Rosenthal suggests excluding such “unphysical” descriptions, but it remains completely unclear how to construe a suitable notion of unphysicality. Rather, the various debates on conventionality in physics have shown that supposedly unphysical descriptions are often feasible and empirically adequate. Furthermore, opinions about physicality habitually change over the course of history. This difficulty is also resolved in a simple manner by the account of causal probability. Essentially, any change in the formulation of the range conditions has to be compensated by a complementary change in measure in order to stay consistent with the collective conditions and the observed frequencies. Obviously, this option is not available to Rosenthal since he insists on using the Lebesgue measure as probability measure. Note again that the same difficulties which Rosenthal makes explicit are hidden in the conditions of microconstancy and macroperiodicity in Strevens’ and Abrams’ account. Strevens’ response in terms of standard variables was already described in Section 2b and is largely equivalent to Rosenthal’s.

Furthermore, there is no need for approximations or imprecisions in the causal account in contrast with Rosenthal’s definition of probability or the related definitions of microconstancy and macroperiodicity in Strevens’ and Abrams’ accounts (cf. the end of Section 2b). Rather, the probability according to the causal interpretation corresponds *exactly* to the weighted fraction of initial states. Again, this move is possible since the causal account renders

²⁵ Remember that the terms ‘causal interpretation’ and ‘causal probability’ now refer exclusively to the account developed in Section 3b.

probability relative to the measure, but also because the causal construal of independence ensures randomness in the sequence of initial conditions and thus convergence of relative frequencies to the causal probabilities by the law of large numbers.

The price to pay is that probability becomes relative to the essentially epistemic choice of a collective and measure, which thwarts the project of a purely objective probability interpretation in deterministic settings. On the other hand, I don't see why accepting some epistemic aspects in probability is problematic except if one adheres to an overly realist view of science. And again, this very step enables the causal interpretation to cover a wide range of applications from indeterministic probabilities to probabilities of hypotheses as described in Section 6—compared with the rather narrow range of applications of the SRA-approach requiring microconstancy and macroperiodicity.

Of course, phenomena accessible to the method of arbitrary functions can be treated within the causal approach as well. In such cases, the collective conditions and the measure need to be fixed only to the extent that the probability distribution is approximately stable. As an example, consider the throw of a die. The probability distribution does not depend much on the exact instructions for the collective conditions, e.g. concerning the original position of the die, the way it is thrown etc. Generally speaking, the choice of collective conditions and measure is largely irrelevant, if the dynamics of the system is sufficiently complex—a topic that is discussed today mainly in the domain of ergodic theory.

On a deeper level, the introduction of measure-dependence in the causal approach calls for new concepts that are not central to the SRA-approach. First, the measure over input states must be determinable independently of relative frequencies in the causal approach—otherwise we would be thrown back on frequentism. To this purpose, the principle of causal symmetry is introduced in the next Section 4. Second, when the condition of microconstancy is dropped, it cannot be assumed anymore that the occurrence of attributes will be sufficiently random due to slight variations in initial conditions. Therefore, in the causal interpretation randomness has to be established by other means leading to the causal construal of independence proposed in Section 5.

With propensity accounts the causal interpretation shares the broad idea that probabilities arise from circumstances or conditions. But otherwise, there are a number of crucial differences. The first point concerns ontology. If the proponents of propensities were thinking of causal determination, why not call it causation? Why use a rather obscure term like propensity? At least Popper seems to have felt the need to introduce a novel ontological category to account for probabilistic phenomena. In later years, he considered causation to be a special case of propensities, namely when the propensity equals one. By contrast, the causal interpretation takes the opposite approach and aims to situate probability within a general framework of causation. While propensity accounts focus conceptually on dispositions or tendencies and rather casually remark upon the parallel with causation, the interpretation proposed in this essay starts with a detailed and specific concept of causation and examines how probability fits into the picture.

On a more methodological level, causal probability is relative not only to the collective conditions but—unlike propensities—also to the measure over initial states. Relatedly, propensity approaches are often silent on the question how exactly the circumstances determine the probabilities. They typically lack the notion of causal symmetry, the ontic version of the principle of indifference, and the causal construal of probabilistic independence. With respect to the last issue, the randomness of subsequent events can be considered as implicit in the notion of tendency in propensity accounts.

Finally, the range of application is usually assumed to be quite restricted for propensities in that they are closely linked with an indeterministic metaphysics—raising the question how the propensity account can be applied in deterministic settings at all. Also, since propensities are often framed in a language of tendencies or dispositions this explicitly excludes the formulation of inverse probabilities, i.e. evidential probabilities or the probabilities of hypotheses (for an elaboration of this criticism, cp. Humphreys 1985). While influential propensity theorists like Popper have argued that inductive concepts like confirmation are not explicable in terms of probabilities at all, the causal interpretation explicitly establishes the link with an inductive framework. Part of the project of a causal interpretation is to show how the basic idea that probabilities arise from circumstances can be extended to epistemic probabilities like the probabilities of hypotheses (cp. Sec. 6).

4. Causal symmetries and the principle of causal symmetry

4a. Causal symmetries

In one article, Rosenthal describes as the “main problem of the range approach” (2010, 81) that it inherits the circularity of the classical approach in that the measure itself requires a justification in terms of probabilities, i.e. probabilities of initial conditions. It might seem that these probabilistic weights of the input states could only be justified on the basis of relative frequencies, which essentially would throw us back on a frequentist account of probability. For authors like Rosenthal, who argue on the basis of the method of arbitrary functions, the solution is to establish that for certain phenomena, most choices of measure lead to roughly the same probabilities. However, as pointed out towards the end of Section 2b, a number of problems result from this approach. These were resolved in the causal approach by rendering probability relative to the measure over initial states (cf. Section 3b).

To tackle the issue of circularity for the causal approach, I will now argue that given full knowledge of the causal setup, probabilities can always be determined by means of symmetry considerations without taking recourse to relative frequencies. In principle, the symmetries must fix the measure only to the extent that a stable probability distribution results. That symmetries and invariances play a crucial role in the determination of probabilities is of course quite obvious, just think of games of chance or Maxwell’s derivation of the velocity distribution in an ideal gas. Of course, for many phenomena the underlying symmetries may not be fully known, which then requires resorting to relative frequencies as a weaker kind of evidence. Referring to the examples of the previous section, population statistics constitutes a

typical case of a frequentist approach to the measure, while the die is a good example for a symmetry approach.

But how exactly the notion of symmetry must be framed in a probabilistic context is not entirely clear from the relevant literature. Let me therefore define as the most basic, if not yet fully general notion of a *causal symmetry*:

A causal symmetry with respect to a probabilistic phenomenon exists if the probability distribution, as determined by the weighted fractions of input states, is invariant under a permutation²⁶ of the attribute space—corresponding to a mere relabeling of the outcome space while all other characteristics of the probabilistic phenomenon remain unchanged including input space, measure over input space, and the dynamics.

The idea that invariance under reformulations can fix a probability distribution has long been used with respect to epistemic symmetries in belief states, reaching back at least to the work of Bolzano (1837/1972, § 161; see also e.g. Jaynes 2003, Ch. 12; Norton 2007). Above, the same kind of reasoning was employed with respect to objective causal symmetries.

In principle, there are two ways in which a causal symmetry can be established, either by referring to the collective conditions or by referring to the dynamics $S \xrightarrow{c} O$. Regarding the former case, note first that the labeling of the outcome space in terms of attributes implies a corresponding labeling of the input space if the dynamics $S \xrightarrow{c} O$ is deterministic. Thus, a known invariance of the measure over the input space under a relabeling may already imply the invariance of the probability distribution under the same relabeling. As an example, the measure over the input space may be determined by a known random process, e.g. the initial state of a ball in an energy landscape may be determined by the throw of a die. In the second case, the invariance of the probability distribution results from the dynamics. Indeterministic phenomena are good examples, as are systems with mixing dynamics. Of course, in general an argument for symmetries will refer both to collective conditions and dynamics.

Only causal symmetries—in contrast to symmetries in belief states—imply the truth of counterfactual statements, such as: If trials of a probabilistic phenomenon were carried out with a different labeling, the probability distribution would remain the same, i.e. any event MX according to the old labeling would have the same probability as the event MX according to the new labeling. With respect to the account of eliminative induction sketched in Section 3b, counterfactual invariance is established by showing the irrelevance of a change in circumstances, in this case of the relabeling of the input and outcome space. This emphasizes once more the importance of the strict method of agreement as a rule for determining irrelevance.

The definition of a causal symmetry directly implies a *principle of causal symmetry* as an objective variant of the principle of indifference:

²⁶ Or more generally, any kind of transformation for a continuous attribute space.

*In the case of a causal symmetry regarding the exchange of two attributes, these attributes have equal probability.*²⁷

Admittedly, the principle verges on tautology, given the previous definition of a causal symmetry. However, the crucial point is that causal symmetries can often be established non-probabilistically, e.g. on the basis of mechanical principles.

As a simple example, consider the fair throw of a fair die. The attribute-space consists in the numbers 1 to 6, located on the different sides of the die. Now, a well-established physical symmetry exists that the numbers on the sides can be permuted in arbitrary ways without changing the probability distribution. Given the principle of causal symmetry, it follows immediately that all attributes must have the same probability 1/6. It is straightforward to apply this type of reasoning to more complex geometrical structures, e.g. a triangular prism with three congruent rectangular sides and two congruent equilateral triangles. Clearly, one can deduce from the corresponding symmetry transformations of the attribute space, that the triangles and the rectangles all have the same probabilities respectively, while not much can be said about the relative probability between rectangles and triangles, except of course that they must add up to one.

The notion of causal symmetry can be extended to more complex transformations of the attribute space including attributes with different probabilities. Such transformations consist in a permutation of the attributes while taking into account the weighted fractions of input states leading to the different attributes. Let $\{M\} = \{M_1, M_2, \dots, M_n\}$ be the attribute space, with $P(M_i)$ denoting the probabilities given by the weighted fractions of input states leading to attributes M_i . Furthermore, let

$\{M'\} = \{M'_1, M'_2, \dots, M'_n\} = T(\{M\}) = \{M_{T(1)}, M_{T(2)}, \dots, M_{T(n)}\}$ be the relabeled attribute space, where $T()$ denotes a permutation of the original attribute space $\{M\}$. Let $P'(M'_i)$ denote the probability of attribute M'_i . Under these circumstances, we can define:

A generalized causal symmetry with respect to a probabilistic phenomenon exists, if for the probability distribution of the permuted attribute space $\{M'\}$ we have:

*$P'(M'_i) = P'(M_{T(i)}) = P(M_{T(i)}) * w(M_i \rightarrow M_{T(i)}) = P(M_i)$, where $w(M_i \rightarrow M_j)$ denotes the ratio of weighted fraction of input states leading to attribute M_i to weighted fraction of input states leading to attribute M_j .*

To avoid circularity, the relative weights $w(M_i \rightarrow M_j)$ should again be established non-probabilistically, e.g. by means of mechanical principles or by causal irrelevance arguments. A corresponding principle of indifference results:

*In case of a generalized causal symmetry, we have: $P(M_{T(i)}) * w(M_i \rightarrow M_{T(i)}) = P(M_i)$.*

²⁷ Note that any permutation can be reconstructed from a sequence of exchanges of attributes. In the case of a continuous attribute distribution and invariance under a certain transformation, the principle of causal symmetry states that an attribute has the same probability as the attribute that it is mapped on.

Obviously, the simpler version of a causal symmetry formulated at the beginning of this section results if $w = 1$. Again a generalization to continuous attribute distributions and their invariance under certain transformations is straight-forward.

Consider as an example of a generalized symmetry a die that is labelled ‘1’ on one side and ‘6’ on all other five sides. The attribute space is $\{M\} = \{1,6\}$ with $\{P\} = \{\frac{1}{6}, \frac{5}{6}\}$. If the attributes are exchanged $\{M'\} = \{6,1\}$ we can calculate as expected $P'(6) = P(6) * w(1 \rightarrow 6) = \frac{5}{6} * \frac{1}{5} = P(1)$ and $P'(1) = P(1) * w(6 \rightarrow 1) = \frac{1}{6} * 5 = P(6)$. Of course, the tricky part is to non-probabilistically establish the causal symmetry and to non-probabilistically determine the relative weight of the attributes $w()$. In the described case of a die, this is rather simple, since the mechanical symmetry with respect to the six sides is fairly obvious, but certainly most applications will be more complex than that.

Instead of transforming the attribute space one could also introduce a complementary mapping of the input space, which leads to a further rendering of the notion of causal symmetry, for example:

*A causal symmetry with respect to a probabilistic phenomenon exists if there is a measure-preserving mapping of the input space onto a different input space, which is still consistent with the collective conditions, leading to a permutation of the attribute space. The attributes that are mapped onto each other have the same probability.*²⁸

Consider for example the throw of a fair coin with a certain set of input states and a measure. Now, by physical reasoning we know: (i) if for every input state the coin is rotated by exactly 180° , then the attributes after the throw will be exchanged: heads \Leftrightarrow tails; (ii) this mapping of the input space is measure-preserving, since for every throw in the original input space there is a corresponding one in the mapped input space. Of course, the mapped input space is still consistent with the collective conditions for the fair throw of a fair coin. Note that the notion of causal symmetry established with respect to the input space also applies to non-probabilistic phenomena, when a phenomenon is invariant under certain symmetry transformations of the initial-state space. For example, the trajectory of a ball is causally symmetric with respect to rotations of the ball in its initial state, given that the mass distribution of the ball is rotationally symmetric.

Let me stress again that causal symmetries are not epistemic judgments in lack of knowledge, but statements concerning the irrelevance of attribute transformations—or, equivalently, transformations of the input space—for the probability distribution. This underlines the significance of the link with eliminative induction, which provides a framework for judging irrelevance on the basis of indifference making.

4b. Further examples

Let us look at more examples of causal symmetries to show that the notion can be applied widely. An interesting case in point is Maxwell’s derivation of the equilibrium distribution for molecular velocities in an ideal gas from symmetry considerations. Here, the attributes are

²⁸ It is again straight-forward to extend this idea to more complex causal symmetries.

labels corresponding to different velocities $\mathbf{v} = (v_x, v_y, v_z)$ and positions in space $\mathbf{s} = (s_x, s_y, s_z)$. Various symmetry assumptions enter in the derivation (Maxwell 1860; cp. Strevens 2013, Ch. 1): (i) homogeneity in space, i.e. there is a causal symmetry with respect to all measure-preserving transformations (relabeling) of the considered spatial volume. It follows that the probability distribution is independent of the spatial coordinates within the considered container (and zero outside the container); (ii) isotropy, i.e. there is a causal symmetry with respect to all rotations (and reflections at the origin) of the velocity space.

This symmetry implies that all velocities with the same absolute value $\sqrt{|v_x^2 + v_y^2 + v_z^2|}$ have the same probability;²⁹ (iii) independence of the one-dimensional velocity distributions along the three Cartesian axes: $P(\mathbf{v}) = f_x(v_x)f_y(v_y)f_z(v_z) = f(v_x)f(v_y)f(v_z)$. Strictly speaking, only the second equality relies on causal symmetry, the first on probabilistic independence.³⁰ As elaborated in Section 5b, probabilistic independence can be established by showing the irrelevance of one attribute distribution for the other. For the sake of simplicity, let us assume just two dimensions x and y. A condition for irrelevance is that the probability $f_y(v_y)$ for any v_y has no influence on the probability $f_x(v_x)$ for any v_x . This holds, since in equilibrium the number of collisions with v_y for one of the particles before the collision and v_x for one of the particles after the collision should be equal to the number of collisions with v_x for one of the particles before the collision and v_y for one of the particles after the collision. Due to this relation, which follows from the constancy of the distribution in equilibrium and from symmetry considerations, changing $f_x(v_x)$ has no influence on $f_y(v_y)$ and vice versa. That the probability distribution is the same $f(\cdot)$ for all coordinates again follows from isotropy. Somewhat surprisingly, these relatively weak conditions (i)-(iii) already hint at the correct probability distribution.

Another causal symmetry is evoked in a later derivation of the equilibrium velocity distribution by Maxwell (1867, 63). In equilibrium one should have the following equality for the probability distributions before and after collisions between two particles: $P(\mathbf{v}_1)P(\mathbf{v}_2) = P(\mathbf{v}'_1)P(\mathbf{v}'_2)$ under the assumption that momentum and kinetic energy is conserved, e.g. $v_1^2 + v_2^2 = v'^2_1 + v'^2_2$ and $\mathbf{v}_1 + \mathbf{v}_2 = \mathbf{v}'_1 + \mathbf{v}'_2$ if all particle masses are the same. Here, primed quantities refer to the velocities after the collision and unprimed before the collision. Again, the relation is not justified by frequency data but by physical reasoning. In fact, it essentially follows from the definition of equilibrium, i.e. the requirement that collisions between particles shall not change the probability distribution: “When the number of pairs of molecules which change their velocities from $[\mathbf{v}_1, \mathbf{v}_2]$ to $[\mathbf{v}'_1, \mathbf{v}'_2]$ is equal to the number which change from $[\mathbf{v}'_1, \mathbf{v}'_2]$ to $[\mathbf{v}_1, \mathbf{v}_2]$, then the final distribution of velocity will be

²⁹ Maxwell argues: “the directions of the coordinates are perfectly arbitrary, and therefore [the probability] must depend on the distance from the origin alone” (Maxwell 1860, 153). This reasoning is criticized by Strevens (2013, 14) on the grounds that Maxwell’s remark supposedly holds for any probability distribution over velocities, which would be an absurd consequence. However, if one understands ‘arbitrary’ in the sense that the choice of coordinates is irrelevant for the probability distribution, then Maxwell’s reasoning is basically correct, evoking a causal symmetry as we had defined it in the previous section.

³⁰ As pointed out by Strevens (2013, 14), Maxwell’s own reasoning in this regard is not entirely convincing, although Maxwell does appeal to independence: “the existence of velocity x does not in any way affect the velocities y or z, since these are all at right angles to each other and independent” (1860, 153).

obtained, which will not be altered by subsequent exchanges.”³¹ (Maxwell 1867, 63) The equality $P(\mathbf{v}_1)P(\mathbf{v}_2) = P(\mathbf{v}_1')P(\mathbf{v}_2')$ can be interpreted as a generalized causal symmetry with respect to transformations of the attribute space $\mathbf{v}_1 \leftrightarrow \mathbf{v}_1'$. It yields direct access to the relative measure $w(\mathbf{v}_1 \rightarrow \mathbf{v}_1') = \frac{P(\mathbf{v}_2')}{P(\mathbf{v}_2)}$. Since supposedly the Maxwell distribution is the only plausible function satisfying the equality, the argument allows establishing this distribution non-probabilistically by appeal to physical symmetries.

A further notable example of causal symmetries concerns the ubiquitous binomial distribution for the calculation of k successes in n trials of an event with probability p : $P_{n,p}(k) = \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$. For the sake of simplicity let us focus on the special case $p = 1/2$. A physical process that generates the corresponding distribution is the Galton board. The essential mechanical symmetry of the Galton board is that at each pin there is no difference between a ball going right or left. Therefore, there is a causal symmetry for each pin i that the probability distribution will not change if one exchanges the labels left l and right r . It follows from the principle of causal symmetry for all i : $P(l|i) = P(r|i) = 1/2$. Based on this insight, the distribution of balls at each level n of the Galton board can be calculated in a purely combinatorial manner by tracing the possible trajectories of the balls through the board. The resulting recursive formula denotes a rather complex causal symmetry that allows to completely determine the binomial distribution at each level $P_n(k) = \frac{1}{2} [P_{n-1}(k-1) + P_{n-1}(k)]$ with $P_0(0) = 1$, $P_n(-1) = P_n(n+1) = 0$. Let me stress again that in deriving the probability distribution for the Galton board we need not make reference to any frequency data whatsoever.

Note, that the mentioned complex symmetry does not immediately fit into the framework described in the previous Section 4a, since the recursive formula relates distributions for different levels n . But it is straight-forward to reformulate it in a way that it fits with the form of generalized causal symmetries $P(M_{T(i)}) * w(M_i \rightarrow M_{T(i)}) = P(M_i)$. Special cases follow directly from further mechanical symmetries of the physical set-up, e.g. $P_n(k) = P_n(n-k)$. A generalization to $p \neq \frac{1}{2}$ is also straight-forward if one can establish a causal symmetry $P(l|i)p = P(r|i)(1-p)$.

The discussion of the binomial distribution directly leads to another important topic, error statistics, which can also be addressed in terms of causal symmetries. Think of each level of pins in the Galton board as a minute contribution to an overall error. According to the central limit theorem, the distribution of this overall error will converge to a normal distribution for large n . More generally, for independent, identically distributed random variables X_1, X_2, \dots, X_n with mean μ and finite nonzero variance σ^2 , the sum $S_n = \sum_{i=1}^n X_i$ will be normally distributed in the limit of large n : $\lim_{n \rightarrow \infty} P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x\right) = \Phi(x)$, where $\Phi(x)$ denotes the cumulative standard normal distribution.³²

³¹ The same relation was used above when arguing for mutual independence of the one-dimensional velocity distributions.

³² Under certain conditions, the requirement of identical distribution for the X can be relaxed.

On the basis of the central limit theorem, one can argue that the sum of a large number of measurements of a certain quantity will be normally distributed, if the measurements take place under identical circumstances and thus the error distribution is always the same. Furthermore, if the error distribution is known to be symmetric around the actual value of the examined quantity, i.e. if one knows that deviations in positive direction occur just as often as those in negative direction, the resulting normal distribution will be centered around the true value. Error statistics thus provides further examples, where arguments concerning causal symmetries that are established non-probabilistically lead to knowledge about probability distributions.

To conclude, let me stress again that the reasoning in these examples does not rely on an epistemic principle of indifference but rather on an objective principle of causal symmetry. Causal symmetries do not refer to lack of knowledge, but follow from the irrelevance of certain transformations of the attribute space for the probability distribution.

4c. The principle of causal symmetry

In Section 4a, I defined the notion of causal symmetry and based on it a principle of causal symmetry as an objective version of the principle of indifference. In its simplest form the principle of causal symmetry states that given a causal symmetry one should ascribe equal probabilities to the corresponding attributes.

How does the epistemic version of the principle of indifference fit into the picture, i.e. the principle of insufficient reason that we should ascribe equal probability when our knowledge about a process does not favor one or the other outcome? Note that there seem to be clear-cut examples, where this epistemic version is employed, for example in Laplace's treatment of the loaded coin: In lack of evidence regarding the way in which the coin is loaded, so the reasoning goes, we should ascribe equal probability to both sides (cp. Section 2a).

Several authors like Cournot or Strevens suggest grounding the distinction between epistemic and ontic probabilities on whether they have been established by an epistemic or an objective version of the principle of indifference, respectively. By contrast, I now argue that apparent applications of the principle of insufficient reason can be reduced to the principle of causal symmetry whenever the resulting probabilities are predictive. The key idea lies in constructing an adequate collective so that the principle of causal symmetry can be applied.

As an example, assume that we know to which extent a coin is loaded, say $p=2/3$, but do not know in which direction. As mentioned, it seems a straight-forward application of the principle of insufficient reason, when one ascribes probability $1/2$ to both heads and tails before the first throw. However, we can also construe an adequate collective to subsume the reasoning under the principle of causal symmetry. The collective conditions should include the premise that the coin is loaded, while the measure ascribes equal weight to both cases $p(\text{heads})=1/3$ and $p(\text{heads})=2/3$. The set-up corresponds to a probabilistic phenomenon, where we are given two coins that are loaded in opposite ways, randomly pick one of them, and throw it. With respect to this collective and measure, the probability $1/2$ for both heads and tails is predictive in terms of limiting frequencies.

When *we know what we don't know* in terms of causal influences on the probability distribution, one can always proceed in this manner, i.e. construct a collective and a measure that account for the lack of knowledge and determine the corresponding probability distribution which is actually predictive in terms of limiting frequencies with respect to the specified collective. Of course, lack of knowledge can come in different degrees. For example, it might be the case that we are only given a probability distribution for the extent to which the coin is loaded. But again, this knowledge already determines the measure and thus an appropriate collective.

Apparently, there are two types of situations, (i) when the collective refers to actual possibilities and (ii) when the range of possibilities is only imagined (cp. also Section 6b). As an example, the two coins that are loaded in different directions could both really exist, e.g. lie on a table before us. Or, there could be just a single coin of which we do not know in which direction it is loaded. In the latter case, the process of randomly choosing between two coins is just imaginary, but a collective and measure must be assigned just as well to avoid contradictions. One might be tempted to ground the distinction between the epistemic principle of insufficient reason and the ontic principle of causal symmetry on this difference between an actual and an imagined collective. But note that conventionally the principle of insufficient reason does not require constructing a causal collective. Also, the mentioned distinction is certainly not sharp but rather blurry, since clearly it is somewhat contextual whether one considers a collective actual or imagined. In any case, the distinction cannot serve to establish a substantial difference between epistemic and ontic probabilities.

Are there applications of the principle of insufficient reason that cannot be reduced to the principle of causal symmetry? Notably, these must be instances where one cannot construct a corresponding collective and measure. In other words, *we do not know what we don't know* in terms of causal influences on the probability distribution. But if collective and measure are underdetermined then we are immediately confronted with Bertrand-type paradoxes. Consider the notorious example concerning the probabilities of different colors, e.g. red, blue, and green. Do red and non-red have the same probability according to the principle of insufficient reason? That cannot be since it would be incompatible with the analogous case that blue and non-blue have the same probability. According to the perspective of this essay, such contradictions arise because the causal context is not specified in terms of collective conditions, range conditions and measure insofar as they are relevant to the probability distribution of attributes. Without the causal context, the principle of indifference leads to contradictions and thus cannot be meaningfully applied.

Thus, Bertrand-type paradoxes are resolved by making probabilities relative to a collective and a measure over input space, i.e. by the requirement that the causal set-up is sufficiently specified. Consider another classic example dating back to Joseph Bertrand himself (1889, 4-5): What is the probability that the length of a random chord in a circle is shorter than the side of an equilateral triangle inscribed in the same circle? Bertrand points out that there are various incompatible answers depending on which measure one chooses, e.g. equal measure for the distance of the middle of the chord to the center of the circle, equal measure for the angle between chord and the corresponding tangent to the circle, or equal measure for the surface element into which the middle of the chord falls. Again, the ambiguity is resolved by

sufficiently specifying the causal process that determines the location of the chord and thereby the measure, e.g. the way a stick is dropped on a circle drawn on the floor.

When the causal context is sufficiently specified in terms of collective and measure, then the corresponding probabilities are automatically predictive about the respective probabilistic phenomenon. Also, under such circumstances, every supposed application of the epistemic principle of insufficient reason can be rendered as an application of the principle of causal symmetry.³³ By contrast, probabilities resulting from applications of the principle of insufficient reason that are not reducible to the principle of causal symmetry with respect to a postulated collective are in general not predictive.

Note finally that the principle of causal symmetry is not affected by another standard objection against the principle of insufficient reason that it supposedly derives something from nothing, namely probabilities from ignorance. Rather, the principle of causal symmetry presupposes considerable knowledge in terms of causal circumstances in order to establish probabilities that are predictive for a specific probabilistic phenomenon. Henceforth, we suggest excluding from the theory of probability all cases where the relevant context in terms of collective and measure is not specified and therefore predictiveness cannot be guaranteed.

5. Causal independence

5a. Randomness

Richard von Mises once presented an interesting example involving posts of different sizes along a road, one large always followed by four small ones. Since the sequence of posts is entirely regular, one would not intuitively consider this a probabilistic phenomenon, at least without a random process determining the location on the road. In fact, von Mises had used this very example to argue that probabilistic phenomena always presuppose a certain irregularity or randomness in the sequence of attributes. Indeed, irregularity ('Regellosigkeit') even constitutes one of the fundamental axioms in von Mises' probability theory. Of course, there exists an obvious reason for this since important theorems in probability theory like the law of large numbers or the central limit theorem presuppose randomness by requiring independence of trials.

Basically, there are two ways, how randomness of an attribute sequence could be determined, either by referring to the dynamics of a process or by establishing randomness in the sequence of input states leading to the various outcomes. A clear-cut example of the first option regards indeterministic phenomena, where the randomness constitutes a law-like feature of the dynamics. An example of the second option are systems in which the input state is itself determined by a random process, e.g. when one is randomly located on the road in the discussed example. An obvious shortcoming of the second option is the threat of circularity, when randomness in outcomes requires randomness in input states. But the more complex the dynamics, the less the actual choice of input states matters.

³³ In this respect, the viewpoint of this essay resembles the position of North (2010), who also denies that there exist distinct objective and epistemic versions of the principle of indifference.

In the past, randomness has mostly been defined with respect to certain properties in the sequence of attributes. Von Mises' discussion of irregularity and Kolmogorov's work on algorithmic complexity are just two examples in this respect. By contrast, randomness can also be associated with independence of trials: *If subsequent trials are independent, then the sequence of outcomes will be random.*³⁴ I will take up this thread in the following.

5b. Independence

The causal approach can throw some light on the notion of independence—an issue that has been called “one of the most important problems in the philosophy of the natural sciences”³⁵ by Kolmogorov. In a recent paper, Strevens essentially concurs and adds that the “matter has, however, received relatively little attention in the literature” (forthcoming, 3). The notion of independence is a major issue in the controversy between subjectivist and objectivist readings of probability. For example, Bruno de Finetti, as a main proponent of subjectivism, aimed to eliminate the essentially objectivist concept of independence altogether and to replace it with exchangeability. In the following, a causal construal of the notion of independence will be sketched deriving independence from causal irrelevance.

For further discussion, it is helpful to distinguish two notions of independence, (i) the independence of consecutive trials of the same probabilistic phenomenon and (ii) independence of random variables associated with different probabilistic phenomena. Roughly speaking, independence of two variables A and B means that (a) one outcome does not affect the other $P(A|B)=P(A)$ or, equivalently from a mathematical point of view, (b) that the corresponding probabilities factorize $P(A,B)=P(A)P(B)$.³⁶ Independence is often defined in terms of such factorization, for example by Kolmogorov (1956, §5). But certainly this does not solve the difficult methodological question how to determine independence in the real world. Why, for example, are two consecutive draws from an urn generally considered independent in case of replacement and otherwise not?

Let us take up a wide-spread intuition and relate independence to irrelevance. In Section 3b, I have argued for a link between eliminative induction and the notion of causal probability. Now, eliminative induction also provides a framework for determining irrelevance in the sense of difference-making with respect to background conditions. Regarding the first notion of independence (i) consider two trials with the same collective conditions and the same measure. A sufficient criterion for probabilistic independence is:

³⁴ Compare the insightful discussion of Poincaré concerning the various ways chance processes can come about (1914, Sec. I.4): (i) if a small cause has large effects; (ii) in the case of complex causes; (iii) if two supposedly unrelated processes interact.

³⁵ “one of the most important problems in the philosophy of the natural sciences is—in addition to the well-known one regarding the essence of the concept of probability itself—to make precise the premises which would make it possible to regard any given real events as independent. This question, however, is beyond the scope of this book.” (Kolmogorov 1956, 9)

³⁶ Note that this covers also the first notion of independence (i), if one interprets the consecutive trials as different probabilistic phenomena.

Two trials are probabilistically independent, if the realization of input states in one trial is irrelevant³⁷ in the sense of difference-making for the probability distribution of input states in the other trial.

Whatever input state is realized in one trial, must not affect the probability distribution of input states in the other trial.³⁸ For example, this establishes the independence of subsequent draws from an urn with replacement or the independence of subsequent throws of a die. In both cases, independence follows from the fact that the causal process determining input states in one instance is irrelevant (with respect to the background of collective conditions) to the causal process determining input states in the other instance. For example, the manner of shaking a die in one instance has no influence on the manner of shaking the die in the other instance. Note that irrelevance again establishes the truth of a counterfactual: If another input state had been realized in the first trial this would not have affected the process determining the input state in the second trial.

The independence of random variables (ii) concerns different probabilistic phenomena that can have different collective conditions and measures over input states. Each random variable is associated with a specific probabilistic phenomenon. A sufficient criterion is:

Two random variables are probabilistically independent, if the realization of input states in one probabilistic phenomenon is irrelevant³⁹ in the sense of difference-making for the probability distribution of input states in the other probabilistic phenomenon.

This criterion broadly stands in the tradition of definition (a) for independence, but it also differs in important respects. Most importantly, it makes reference not to the attribute distribution but to the distribution of input states. Thus, the evaluation of the criterion is more intuitive since it makes explicit reference to the processes that are causally responsible for the probability distributions of attributes. As an example the throw of a coin and the probability of rain tomorrow are independent, because there is no causal connection between the corresponding processes determining the input states. On the other hand, the probability of smoking and the probability of getting lung cancer are in general not independent in an individual, because there is a plausible causal influence from the range conditions of smoking to those of getting lung cancer.

Note again that with respect to the conventional definition of independence the criteria given above are only sufficient but not necessary. As an example, consider two consecutive draws of a ball. The first ball is drawn arbitrarily from one of two urns B and W both of which have the same ratio of black and white balls. The second draw depends on the result of the first draw. If the ball is black, the next one is drawn from urn B, otherwise from urn W. Now, even though there is causal relevance of the input states in the first draw for those of the second draw, the draws are still independent in the conventional sense: for the attribute black/white in the second draw the attribute of the first draw does not matter. The trick is of course that

³⁷ i.e. irrelevant with respect to a causal background fixed by the collective conditions.

³⁸ Note that independence is not a directed concept, i.e. independence of one trial from another automatically implies independence of the latter from the former as well.

³⁹ i.e. irrelevant with respect to a causal background fixed by the collective conditions.

while there is causal dependence of the input states, this has no influence on the probability distribution in the second draw.

Thus, one could conceptually distinguish probabilistic independence as framed above in terms of irrelevance of the input states from the conventional concept of probabilistic independence referring to the irrelevance of attributes. Of course, the former implies the latter—simply because the attributes are defined on the outcome space which is determined by the input space via the causal mapping $S \xrightarrow{c} O$. A sufficient and necessary criterion for independence in the conventional sense is:

Two trials are probabilistically independent iff the realization of attributes in one trial is irrelevant⁴⁰ in the sense of difference-making for the probability distribution of attributes in the other trial.

For instance, if the causal mapping $S \xrightarrow{c} O$ is sufficiently complex, the first version in terms of input states will generally be too restrictive. In such cases, there may be substantial dependence of subsequent input states, but independence in the sequence of attributes could for example be realized, if microconstancy holds and there is small, but sufficient variability in the initial conditions.

Equally:

Two random variables are probabilistically independent in the conventional sense iff the realization of attributes in one probabilistic phenomenon is irrelevant⁴¹ in the sense of difference-making for the probability distribution of attributes in the other probabilistic phenomenon.

Essentially, this is only the familiar requirement $P(A|B)=P(A)$, while specifying that the criterion is to be understood in terms of irrelevance according to eliminative induction. An example was discussed in Section 4b concerning the mutual independence of velocity distributions along different coordinate axes in an ideal gas at equilibrium.

In summary, we have suggested how probabilistic independence could be derived from causal irrelevance of probabilistic phenomena as determined by eliminative induction. Of course, these few sketchy ideas cannot account for the enormous complexity of the notion.

6. Ontic and epistemic probabilities

6a. Single-case probabilities and indeterminism

Indeterministic phenomena can easily be integrated into the suggested framework of causal probability. For a fully indeterministic phenomenon, there are no hidden variables, i.e. no range conditions that determine outcome and attribute. More exactly, there is only one input state and the measure over input space thus becomes trivial. With respect to the terminology

⁴⁰ i.e. irrelevant with respect to a causal background fixed by the collective conditions.

⁴¹ i.e. irrelevant with respect to a causal background fixed by the collective conditions.

introduced in Section 3b, there are no measure conditions and the collective conditions consist only of set-up conditions, which by means of the indeterministic dynamics $S \xrightarrow{c} O$ fix a measure over the outcome space and thus the probability distribution for the attributes. Adjusting the definition of probability of Section 3b accordingly, that it makes reference to a probability measure over the outcome space instead of the input space, carries no further difficulty.

The orthodox interpretation of quantum mechanics yields a prime example. Via the Schrödinger equation, the collective conditions determine the wave function and thereby the probability distribution for certain attributes like position or momentum. The orthodox interpretation explicitly excludes range conditions which would correspond to hidden variables rendering the phenomenon deterministic.

These remarks can also help to clarify the role for single-case probabilities according to the perspective of this essay. In principle, there are no probabilities without collective and measure. However, fully indeterministic events could be interpreted in terms of single-case probabilities, since for these there exists a natural choice of collective conditions and measure. Note further that according to the causal approach of this essay one can speak of the probability of an event, even though the corresponding probabilistic phenomenon may have occurred only once. As long as one has access to the measure over input space, the phenomenon need not even be repeatable. This distinguishes the causal approach from the naïve frequency view which obviously has to rely on a sufficient number of instantiations.

6b. Epistemic and ontic probabilities

The discussion of indeterminism in the previous section directly leads to one of the basic themes in the debate on interpreting probability, namely the distinction between epistemic and ontic probabilities. As emphasized before, unlike the SRA-approach, the causal framework delineated in this article is meant to extend to cases of indeterminism and also to epistemic probabilities such as probabilities of hypotheses. In fact, causal probability is intended to cover all applications of the probability axioms in which probability is predictive.

The definition from Section 3b allows identifying different types of probabilities along the ontic-epistemic spectrum. (i) Purely ontic probabilities are those for which a specific collective is singled out by the statistical event. The typical example concerns indeterminism as discussed in Section 6a, e.g. the decay of a radioactive atom according to the orthodox interpretation of quantum mechanics.

(ii) When the event does not single out collective conditions and measure, there will automatically be an epistemic element in the choice of these quantities. First, there remains some leeway, which causal circumstances to consider as collective conditions and which as range conditions, usually implying a change in probabilities. Second, the probability measure over input states will not entirely be fixed objectively. To some extent, it remains a matter of choice, how often an input state will appear in a probabilistic phenomenon. In principle, these epistemic dimensions also exist for the deterministic probabilities established by the method of arbitrary functions, if somewhat less pronounced.

(iii) A further epistemic element concerns the distinction between a situation, where the possibilities determined by the collective conditions are realized in the world, and situations, where the possibilities are to various extent only imagined or postulated. For example, does one actually choose between two coins that are loaded in different ways—or is there only one coin and is the ensemble of two coins just imagined as a subjective range of alternatives? These two cases roughly correspond to the distinction between an objective and an epistemic reading of the principle of indifference, as introduced in Section 4c. In the first case, the causal symmetries on which probabilities are based really exist in the world. In the second case they are imagined: ‘if the collective were such and such, then the following probability distribution would be predictive.’

As noted before, the distinction is not sharp and depends considerably on context. But of course, when the various possibilities are not realized in the world, there is considerable flexibility how to construct collective and measure—corresponding to a more pronounced subjective element in the assignment of probabilities.

In the following, I will discuss two further variants of epistemic probabilities, first concerning predictions that rely on symptoms and not on causes and second probabilities of hypotheses.

6c. Probabilities from causal symptoms

Sometimes, the input space is parametrized not in terms of causes of the probabilistic phenomenon, but rather in terms of symptoms or proxy variables that are somehow causally related. A typical example concerns the correlation between barometer and weather. One can quite reliably predict the weather by referring to a barometer reading, but of course the barometer reading is not a cause of the weather. Rather, air pressure is a common cause that influences both barometer and weather. Since air pressure is not easily accessible epistemically, one might be tempted to postulate a probabilistic phenomenon that has as input space the barometer reading and as output space a certain parametrization of the weather. While in practice such probabilities predicting from symptoms or proxies of common causes are wide-spread, let us briefly examine if they are consistent from the view of causal probability.

Formally, we have an outcome space O , a space spanned by the parametrization of the symptoms I , and an unknown input space S that causally determines the outcome space. In the example above, O would be the weather, I would be the barometer reading, and S would be spanned by some microparameters determining the weather, including air pressure. Two situations need to be distinguished: (i) the symptoms I are fully determined by S ; (ii) there are other causes of I that are not in S .

In the first case, probabilities from symptoms easily fit into the framework of causal probability in the following manner. For the sake of simplicity, assume that to any S can be attributed an I . The symptoms I can then be considered as labels of the input space and thus as a reparametrization of the input space, which allows to establish a probability distribution over the attributes based on the symptoms. Note that the mapping $I \rightarrow O$ will in general not be fully deterministic, i.e. the same value of I can lead to different values of O .

By contrast, such a probability distribution does not exist in the second case, because there are other unrelated causes for I. For example, someone may mechanically interfere with the barometer reading or the spring in the barometer may break for various reasons. If such external causes are possible, then a probability distribution for the attributes based on symptoms cannot be given. The situation can only be resolved, if one includes in the parametrization of the input space S all possible external causes of I and if one knows the probability measure over those causes. In that case, we can again interpret the symptoms I as a reparametrization of the extended input space and a meaningful probability distribution results for the attributes.

In summary, probabilities from symptoms are only meaningful if they can in principle be reduced to causal probabilities as defined in Section 3b.

6d. Probabilities of causal hypotheses

Thus far, we have treated probabilities of events or types of events as determined by their causal circumstances. But the inductive framework of Section 3b can also cover inverse probabilities, i.e. probabilities of hypotheses regarding possible causes generating the given evidence. The reason is that the eliminative logic underlying causal probability works in both directions—from given causes to possible effects and from given effects to hypotheses about causes.

Consider again a probabilistic phenomenon determined by certain collective conditions, an input space, a measure W over the input space and a causal mapping from input space to outcome space. When determining the probability of hypotheses, a labelling of the input states must be introduced, which allots these to the different hypotheses H_1, \dots, H_N . This labelling must be mutually exclusive and must cover the whole input space. If the causal mapping is bijective⁴², a corresponding labelling of the outcome space results. The causal mapping also determines a measure W_O over the outcome space from the measure over the input space. Relevant evidence leading to an adjustment of the probabilities of the various hypotheses can concern the input space and the outcome space. We can now define:

The probability of a causal hypothesis HX , combining a set of input states of the probabilistic phenomenon P , is given by the fraction of input states weighted with measure W carrying the label HX or, equivalently, by the fraction of outcome states weighted with measure W_O carrying the label HX .⁴³

Let us look at the Monty Hall problem as a simple example for probabilities of causal hypotheses generating a given evidence. In a quiz show, a candidate is presented with three doors A, B, C, behind one of which is a car, behind the two others there are goats. The candidate chooses one of the doors, e.g. A. At the beginning, the evidence conveyed by the quizmaster does not favor any of the hypotheses H_A, H_B, H_C that the car is behind the respective door. In other words, there is a causal symmetry in the set-up of the game with

⁴² It may also be indeterministic.

⁴³ Note that the probabilities of hypotheses can be interpreted in terms of probabilities of events, when it is possible to look up which of the hypotheses is actually realized in the world. For example, in the Monty Hall problem discussed below, the corresponding event would consist in opening all doors to verify where the car is.

respect to permutations of the doors A, B, C. Consequently, the labels are equally distributed in both weighted input and weighted outcome space, resulting in equal probability for all three hypotheses.

Now, the quizmaster opens a door, e.g. C, of which he knows that there is a goat behind it and which is not the one chosen by the player. Thereby, new information E is conveyed—which can be accounted for in terms of an additional collective condition. In light of this new condition, the input states which are incompatible with E have to be erased. In particular, all input states of hypothesis HC have to be eliminated, because the truth of HC is incompatible with the evidence. Furthermore, half of the input states of hypothesis HA have to be eliminated, namely those, in which the quizmaster would have opened door B. By contrast, none of the input states of HB are deleted because all of them already imply that the quizmaster opens door C. This leads to the familiar result that in light of the new evidence we have $P(HA)=1/3$ and $P(HB)=2/3$.

Obviously, this result can also be calculated via Bayes' Theorem: $P(HX|E) = \frac{P(E|HX)P(HX)}{\sum_{i=1}^N P(E|Hi)P(Hi)}$. The quantities on the right side refer to the old collective, $P(HX|E)$ on the left side is equivalent to the probability $P(HX)$ relative to the new collective incorporating evidence E. In summary, the change in collective conditions due to novel evidence corresponds to a process of Bayesian updating.

Another example concerns the loaded coin as already discussed in previous sections—except that this time we are not interested in the event of throwing the coin, but in the probability of the two hypotheses H1 and H2 that the coin is loaded $P(\text{heads})=2/3$ or $P(\text{heads})=1/3$, respectively. Before the coin is thrown for the first time, the evidence does not favor any of the hypotheses and therefore both hypotheses have equal probability 1/2 with respect to a suitably constructed collective. After the first throw, the situation ceases to be symmetric since there is now evidence in which way the coin might be loaded. Again, this evidence can be integrated in the collective conditions leading to a change in measure and thus a new probability distribution over the causal hypotheses. For example, if the result is 'head', then all those input states have to be eliminated that would have led to 'tail' in the first throw, i.e. 1/3 of the input states belonging to H1 and 2/3 of the input states belonging to H2. The new probabilities are consequently $P(H1)=2/3$ and $P(H2)=1/3$, which is exactly the result given by Bayes' Theorem. From the causal perspective, Bayesian updating can be interpreted as describing how in light of new evidence, which leads to additional constraints in the collective conditions, the measure over the hypothesis space has to be adapted.

Also in the case of probabilities of hypotheses, the ascription of probabilities is predictive only if one specifies collective and measure, i.e. in particular if one knows the complete set of (mutually exclusive) causal hypotheses and if one knows or assumes a measure over these hypotheses that is determined by the collective conditions. Of course, one also needs to know with which probabilities the different hypotheses lead to various pieces of evidence, i.e. essentially the causal mapping of the input to the outcome space. These requirements delineate a fairly restricted range of application for probabilities of hypotheses—excluding for example several 'standard' applications of subjective Bayesianism like the probabilities of

abstract scientific theories or hypotheses. Since the range of alternatives is not known in these cases, it seems implausible to construct a collective and relatedly the measure remains undetermined. If one requires probabilities to be predictive, the range of hypotheses to which probabilities should be ascribed is thus rather restricted.⁴⁴

We are therefore in the position to assess the plausibility of the various Bayesian programs from the perspective of causal probability. Sometimes, the hypothesis space and the measure are objectively determined by the causal set-up. Consider for example the following experiment with three urns, each containing both black and white balls but in different ratios, e.g. 1:2, 1:1, 2:1, corresponding to three hypotheses. Now, one of these urns is randomly chosen and then balls are drawn with replacement. Given a certain sequence of draws as evidence, e.g. w-w-b, a probability for each of the three hypotheses can be calculated. In this specific situation, an objective Bayesian approach is feasible because all relevant elements are determined by the physical set-up: the hypothesis space, the initial probability measure over the hypothesis space, and the probability of evidence given a certain hypothesis is true.

In other circumstances, we might not be so lucky. We may for example be confronted with limited information about a single urn, e.g. that the colors of the balls are only black and white and that there are no more than five balls in the urn. In this case, the hypothesis space is determined by the set-up but there is flexibility in the choice of measure since the actual process with which the urn was prepared is unknown. In analogy to the discussion in point iii) of Section 6b, the Bayesian can now construct in her mind a collective to which the urn is attributed, e.g. an ensemble in which every ratio of balls has equal prior probability. With respect to such a collective, the posterior probabilities of the various hypotheses can then be calculated taking into account additional evidence. However, the Bayesian might just as well have chosen a different measure over the hypothesis space and would have come up with a different result for the posterior probabilities. There is no contradiction, since strictly speaking the probabilities only hold relative to the respective collective and measure. In cases, where the measure is underdetermined by given knowledge and somewhat arbitrarily construed with respect to an imagined collective, we may plausibly speak of subjective Bayesianism.

Of course, much more should be said how Bayesianism is to be integrated into the framework of causal probability. But the brief discussion above already suggests how the notion of causal probability allows determining the limits of a Bayesian approach.

7. Conclusion

We have proposed in this essay a specific account of causal probability that ties in with recent work on objective probabilities in the tradition of the method of arbitrary functions and with earlier accounts mainly from the 19th century, for example by Cournot, Mill, or von Kries. The causal probability of this essay broadly fits with eliminative induction and the

⁴⁴ An argument in this direction was already given by Popper, who claimed in a *reductio ad absurdum* that given an infinite number of alternatives, the probabilities of scientific theories would always be zero. See also Pietsch (2014) for a different argument against ascribing probabilities to scientific theories or abstract scientific hypotheses.

corresponding difference-making account of causation. Probability is rendered relative to the causal conditions of an event in terms of collective conditions and a measure W over the input space. The proposed notion of probability is the following: *The causal probability of a specific attribute MX of a probabilistic phenomenon P is given by the fraction of input states weighted with the probability measure W leading to attribute MX .*

As a further constraint, we required that one should speak about probabilities only when they are predictive. This delineates the range of application for probabilities both of events and of hypotheses. It also allows for a refined version of the principle of indifference, which was termed principle of causal symmetry and which also covers supposed applications of the epistemic principle of insufficient reason. Note again that the principle of causal symmetry does not fall prey to Bertrand-type ambiguities because it requires that the causal context is sufficiently specified. Regarding the difficult notion of probabilistic independence a suggestion was sketched how to connect it to causal irrelevance based on eliminative induction. The mentioned definition of probability, the notion of causal symmetry, and the causal construal of probabilistic independence should be considered as a coherent package making up causal probability.

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