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Structural Inequality in Collaboration Networks

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5 **Abstract**

6 Recent models of scientific collaboration show that minorities can end
7 up at a disadvantage in bargaining scenarios. However, these models
8 presuppose the existence of social categories. Here, we present a model
9 of scientific collaboration in which inequality arises in the absence of
10 social categories. We assume that all agents are identical except for
11 the position that they occupy in the collaboration network. We show
12 that inequality arises in the absence of social categories. We also show
13 that this is due to the structure of the collaboration network and that
14 similar patterns arise in two real-world collaboration networks.

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26 1 Introduction

27 Science is a social enterprise. For the most part, scientists do not work in
28 isolation but collaborate with others when running experiments, analyzing
29 data, or publishing papers. Scientific collaborations have in fact become
30 more common over the past decades throughout academic disciplines (Melin
31 and Persson, 1996; Henriksen, 2016). On the bright side, collaborations can
32 bring about a host of epistemic and practical goods: collaborations seem to
33 increase research output and impact (Beaver, 2004; Lee and Bozeman, 2005),
34 and may even promote the attainment of truth by allowing researchers to pool
35 resources and expertise (Wray, 2002).

36 But the social dimension of science can also bring about unequal out-
37 comes, as philosophers of science have recently shown. Drawing on results
38 from Bruner (2019) and O’Connor (2017), O’Connor and Bruner (2019) show
39 that minorities can end up at a disadvantage in bargaining models of scien-
40 tific collaboration merely because of their group size. Similar models suggests
41 that minority disadvantage can hinder progress in epistemic communities
42 (Rubin and O’Connor, 2018), and that intersectionality may aggravate the
43 issue (O’Connor et al., 2019).¹

44 Models of inequality in scientific collaboration can be very illuminating:
45 they provide a possible account of how discrimination against minority groups
46 might arise without explicit or implicit bias, or indeed without any difference
47 between groups apart from size. But so far models of inequality in scientific
48 collaboration presuppose the existence of social categories, with agents differ-
49 ing in some arbitrary but visible trait—e.g. race, gender, age, or membership
50 in some other social group. One may therefore be led to conclude that social
51 categories are the main or perhaps the only cause of inequality in epistemic
52 communities. Conversely, it would be a lot more troublesome if inequality
53 could arise in the absence of social categories. Inequality might then persist
54 even if we could somehow erase the divides between distinct social groups.

¹The social dimension of science can lead to outcomes that are undesirable for epistemic reasons as well. For example, community size and connectivity can restrict how quickly scientists converge on the truth (cf. Rosenstock et al., 2017; Zollman, 2007, 2010). When facing a risk-return trade-off in their work, individual scientists can divide cognitive labor in ways that are suboptimal for the community as a whole (Kummerfeld and Zollman, 2015); see also Kitcher (1990) and Weisberg and Muldoon (2009). Other social aspects of research, such as the influence of funding agencies, can bias epistemic communities and steer scientists away from the truth (Weatherall et al., 2020; Holman and Bruner, 2017).

55 Here, we present a model of scientific collaboration in which inequality
56 arises in the absence of social categories. Our model represents a collabora-
57 tion network where scientists must bargain over how much effort to invest in
58 joint projects and how to divide credit for their labor. We then show that
59 some scientists can end up at a disadvantage when all scientists are identical
60 except for the position they occupy in the collaboration network. We also
61 show that this unequal outcome is due to the structure of the collaboration
62 network. Inequality thus emerges in the absence of biases or social categories,
63 although biases and social categories may compound the problem.

64 The paper proceeds as follows. We begin by reviewing previous results in
65 Section 2. We then describe and justify our model in Section 3. In Section
66 4, we report results from computer simulations showing that the structure of
67 collaboration networks can lead to inequality in the absence of social cate-
68 gories. We also show that similar patterns arise in two real-world collabora-
69 tion networks and that different dimensions of inequality can come apart. In
70 Section 5, we discuss how our findings relate to previous work on bargaining
71 models of scientific collaboration. We conclude in Section 6 by considering
72 some limitations of our approach.

73 **2 Previous Models**

74 Recent models of scientific collaboration focus primarily on inequalities that
75 arise due to social categories. There are good reasons for this, as inequality
76 in scientific practice is often linked to social markers. The gender gap is a
77 particularly well-documented case. Female scientists tend to publish fewer
78 papers than male colleagues and are less likely to participate in collaborative
79 research projects (West et al., 2013; Larivière et al., 2013). Female scientists
80 also receive grants less often when funding agencies assess their quality as
81 principal investigators, but not when agencies assess the quality of their
82 research proposals (Witteman et al., 2019). There is further evidence that
83 young female scientists are less likely to be listed as an author in a published
84 paper, despite working more hours in total than male colleagues (Feldon
85 et al., 2017). Similar patterns of discrimination arise with respect to race
86 and ethnicity as well: in many disciplines, members of underrepresented
87 racial and ethnic groups tend to have fewer publications and lower promotion
88 rates (Hopkins et al., 2013; Gabbidon et al., 2004; Abelson et al., 2018).

89 In an effort to understand inequality of this form, previous models of

90 scientific collaboration consider a simple version of the Nash demand game
 91 (Nash, 1950). In this game, two agents decide how to split a resource by
 92 demanding a portion of it. If the sum of their demands is equal to or less
 93 than the total amount available, each agent gets what they demand. If the
 94 sum of their demands exceeds the total amount, each agent gets nothing on
 95 the assumption that the negotiation breaks down when they cannot come to
 96 an agreement. For simplicity, we assume that agents can only make one of
 97 three possible demands: low (*Low*), medium (*Med*), or high (*High*). This is
 98 the mini-Nash demand game (Skyrms, 1996), with payoffs shown in Table 1.

Table 1: **Payoffs in the mini-Nash demand game.** In each cell, the first and second entries represent the payoff to the row and column players. Note that $L < M = 0.5 < H$ and $L + H = 1$.

	<i>Low</i>	<i>Med</i>	<i>High</i>
<i>Low</i>	L, L	$L, 0.5$	L, H
<i>Med</i>	$0.5, L$	$0.5, 0.5$	$0, 0$
<i>High</i>	H, L	$0, 0$	$0, 0$

99 When agents are perfectly rational, any two demands that sum to 1 is a
 100 pure Nash equilibrium of the game. Given any such configuration, neither
 101 agent has an incentive to unilaterally demand a different share of the resource.
 102 For example, there is an equilibrium where both agents demand *Med* and
 103 split the resource evenly. Such equilibria are usually termed “fair”. There
 104 are also mixed Nash equilibria in which agents mix two or all three demands
 105 with some positive probability. For example, there is an equilibrium in which
 106 one agent demands *Low* with probability L/H and the other demands *High*
 107 with probability $1 - L/H$. Such equilibria are usually called “unfair”.

108 Equilibrium results differ when agents are not perfectly rational and in-
 109 stead adjust their strategy via a process of biological or cultural evolution.
 110 Using the replicator dynamic as a model of evolution, Skyrms (1996) shows
 111 that there are only two equilibria in a population of agents playing the mini-
 112 Nash demand game: a symmetric equilibrium with agents who only play
 113 *Med*, and a mixed equilibrium with some agents playing *Low* and others
 114 playing *High*. Both equilibria are stable. But the equilibrium in which
 115 agents play *Low* and *High* is inefficient: when two agents demanding *Low*
 116 meet, each gets a positive payoff but a portion of the resource goes to waste.

117 This inefficient equilibrium can be avoided. If agents differ on the basis
 118 of arbitrary but visible group markers, agents can make their strategy condi-

119 tional on the group membership of others. In this way, agents can coordinate
120 on one of the efficient equilibria (Skyrms and Zollman, 2010). The popula-
121 tion then evolves to either the symmetric equilibrium in which everyone plays
122 *Med*, or the asymmetric equilibrium in which one group demands *High* and
123 the other group demands *Low*. The asymmetric equilibrium is known as a
124 “discriminatory norm”: a self-reinforcing pattern of behavior that puts some
125 at a disadvantage merely because of group membership (Axtell et al., 2001).

126 Interesting outcomes are also possible when the population is divided
127 into groups that have different sizes. Although the symmetric equilibrium is
128 still stable in this case, Bruner (2019) and O’Connor (2017) show that the
129 smaller the minority group is, the more likely the population is to evolve to an
130 equilibrium with the minority demanding *Low* and the majority demanding
131 *High*. Similar results have been observed in experiments where participants
132 play the mini-Nash demand game in groups of different sizes (Mohseni et al.,
133 2019). Under these conditions, the minority is more likely to demand *Low*
134 because the minority encounters the majority more often than the other way
135 around. As a result, the minority is faster to adapt to the demands of the
136 majority. This outcome is the cultural analogue of the Red King effect: when
137 two populations co-evolve, the population that is slower to adapt gains the
138 evolutionary upper hand (Bergstrom and Lachmann, 2003).

139 Bargaining games such as the mini-Nash demand game have a long his-
140 tory as models of resource division (Skyrms, 1996; Binmore, 1998). Recently,
141 the mini-Nash demand game has also been used to model the division of re-
142 sources resulting from scientific collaborations. O’Connor and Bruner (2019),
143 for example, use the mini-Nash demand game to show that members of the
144 minority group can end up at a disadvantage in scientific collaboration sim-
145 ply because of their group size. Rubin and O’Connor (2018) draw on similar
146 models to describe how discrimination can lead to segregation, which de-
147 creases the diversity of collaboration networks and is thus likely to hinder
148 epistemic progress in science.

149 In the next section, we describe a model using the mini-Nash demand
150 game to represent the division of resources resulting from scientific collabo-
151 ration. But there are no social categories in our model. Yet, we show that
152 inequality can arise because of the structure of the social network.

153 3 Model Description

154 The mini-Nash demand game captures important features of scientific collab-
155 orations (Rubin and O’Connor, 2018; O’Connor and Bruner, 2019). Scientists
156 must often decide whether or not to enter a collaboration. If they choose to
157 join the project, they must decide how to divvy up the credit for their joint
158 labor. We therefore take a strategy in the mini-Nash demand game to repre-
159 sent a request for a certain amount of credit resulting from the joint project.
160 One example of how a scientist might claim credit is by requesting to be first
161 author. But there are other ways in which a scientist might claim credit. For
162 example, a scientist might claim credit by explicitly describing their role in
163 an author contribution statement, presenting results from the joint project
164 at a conference, or promoting the project in social media. The *Low* strategy
165 thus corresponds to a case in which a scientist requests a small amount of
166 credit, the *Med* strategy to a case in which a scientist demands a moderate
167 amount of credit, and the *High* strategy to a case in which a scientist de-
168 mands a large amount of credit. We assume throughout that collaborators
169 do enough work to get an output of sufficient quality, thus ensuring that
170 research quality is held constant.

171 Accordingly, the *Low – Low* outcome might correspond to a case in which
172 both scientists evince a certain level of timidity, do not promote the project
173 in social media or do not present it at conferences, and therefore claim only
174 a small amount of credit. In this case, both scientists split the credit evenly
175 but claim a small amount of credit in total so each scientist ends up receiving
176 a low payoff. In the *Med – Med* outcome, both scientists claim a moderate
177 amount of credit—for example, by promoting the project in social media or
178 presenting it at conferences. In this case, scientists again split the credit
179 evenly but each scientist claims a moderate amount of credit and so ends
180 up receiving a moderate payoff. In the *Med – Low* outcome, the scientist
181 playing *Med* claims a moderate amount of credit while the scientist playing
182 *Low* claims a small amount of credit. Thus, the *Med* scientist gets a moderate
183 payoff and the *Low* scientist ends up with a small payoff. In the *High – High*
184 and the *High – Med* outcomes, both scientists claim too much credit for
185 themselves and conflict erupts between them. As a result, the collaboration
186 breaks down and both are left with a payoff of zero.

187 In line with this interpretation of the *Low*, *Med*, and *High* strategies, we
188 use the mini-Nash demand game to represent the division of credit in scientific
189 collaborations. In contrast to these models, however, we assume that there

190 are no social categories. We make this assumption because in some cases
191 inequality in science does not appear to be due to social categories, being
192 rather linked to the structure of the social network. A case in point is the
193 “Matthew effect” (Merton, 1968). The Matthew effect describes how more
194 prominent scientists often get more credit than less prominent ones for work
195 of equal worth. Since the mechanism was first proposed, empirical studies
196 have confirmed that the Matthew effect is pervasive in science. For example,
197 early work shows that inequality in publication counts increases as scientists
198 age, suggesting a cumulative effect over time (Allison and Stewart, 1974;
199 Allison et al., 1982). Recent work indicates that citation counts appear to
200 depend in part on how renowned the author already is (Petersen et al., 2014).
201 In fact, the problem seems to be getting worse (Nielsen and Andersen, 2021).
202 A Matthew effect can also be seen in science funding, with recipients of early-
203 career grants being more likely to win further grants than equally qualified
204 peers (Bol et al., 2018).

205 In light of the evidence that inequality is not always directly due to social
206 categories, we consider how inequality can arise in scientific communities in
207 the absence of social categories. As there are no social categories in our
208 model, we assume that scientists are identical except for the position they
209 occupy in the collaboration network. In particular, we let scientists occupy
210 the N nodes of a graph. Further, we let $e_{ij} = 1$ represent a link between
211 scientists i and j if they collaborate on a joint project and $e_{ij} = 0$ otherwise.
212 Scientist i then plays the mini-Nash demand game with every scientist j such
213 that $e_{ij} = 1$. For simplicity, we assume that every scientist i plays the same
214 strategy with all their collaborators. In each round of interaction, their total
215 payoff is then given by the following expression:

$$\pi_i = \sum_j^N e_{ij} \cdot r_{ij} \quad , \quad (1)$$

216 where r_{ij} is the reward that i gets from interacting with j . The total payoff is
217 thus the sum of rewards that a scientist receives from all their collaborators.²

218 As before, we suppose that scientists receive rewards according to Table
219 1. Since the values of L and H determine how large the gap is between
220 the rewards that *Low* and *High* scientists get, we take these parameters to

²We consider the sum, and not the average, of rewards because it is more natural to think of scientists adding the rewards they receive from joint projects instead of averaging them. But results are the same if we instead take the average reward.

221 represent how “elitist” or “egalitarian” a scientific community is with respect
222 to reward allocation. A large difference between L and H thus represents
223 an elitist community where scientists either get a very low or a very high
224 reward; in contrast, a small difference represents an egalitarian community
225 where scientists mostly get the same reward. Indeed, scientific communities
226 appear to differ in how unequal they are (Han, 2003; Clauset et al., 2015).³

227 To model the structure of the scientific community, we turn to sciento-
228 metric studies on the topology of collaboration networks. Empirical evidence
229 suggests that collaboration networks often have predictable properties, de-
230 spite discipline-specific idiosyncrasies. In particular, collaboration networks
231 tend to have a skewed degree distribution (Newman, 2001, 2004). This is
232 to say that the distribution of the number of collaborators per scientist has
233 a long tail, with collaboration networks displaying a hub-and-spoke archi-
234 tecture in which few scientists (“hubs”) have many collaborators and many
235 scientists (“spokes”) have just a few. More precisely, the degree distribution
236 of collaboration networks has the following form:

$$P(d) \sim d^{-\gamma} \quad , \quad (2)$$

237 where γ controls the shape of the distribution and d is the degree or the
238 number of collaborators per scientist. Networks with a degree distribution of
239 this form are known as “scale-free”. A similar degree distribution is common
240 in other social and biological networks, such as animal societies and gene
241 regulatory networks (Barabási and Oltvai, 2004; Lusseau, 2003).

242 For this reason, we consider here scale-free networks with a power-law
243 degree distribution. Although there are many models of network formation
244 that result in such a distribution, a simple model that is known to gener-
245 ate a power-law degree distribution is the preferential-attachment model due
246 to Barabási and Albert (1999). In this model of network formation, there
247 is initially a small set of interconnected nodes. Nodes are then added to
248 the network and connected to other nodes with probability proportional to
249 the number of connections that existing nodes already have, giving rise to a

³As an anonymous referee points out, some academic communities have a reputation for being especially elitist—e.g. economics. At the same time, economics follows a strict norm of alphabetical author order implying equal contribution in collaborative work. This might be taken to mean that economics is an egalitarian discipline after all. However, it is possible that an alphabetical author order only makes a discipline more elitist: if authors do not disclose their real contribution to a joint project, others must resort to an author’s past reputation or institutional affiliation to infer their real contribution.

250 Matthew effect in network formation. As the network grows, few nodes ac-
 251 cumulate many connections and many nodes acquire only a few. In the limit
 252 of an infinitely large network, the resulting degree distribution converges on
 253 the power law given by equation (2). There are certainly more sophisticated
 254 models of network formation, but the preferential-attachment model is a sim-
 255 ple and widely used one. For comparison, we consider regular networks in
 256 which every node has the same degree d and thus the average degree is also
 257 d . In particular, we consider regular networks with $d = 2$ and $d = 5$. These
 258 regular networks are not realistic but serve as control cases, as the scale-free
 259 networks we analyze have an average degree of about $d = 2$ (see Figure 1).

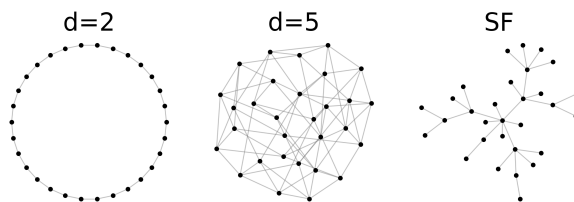


Figure 1: **Network topologies.** *Left:* regular network with $d = 2$. *Center:* regular network with $d = 5$. *Right:* scale-free network given by the preferential-attachment model described in Barabási and Albert (1999) with one initial node. Shown are networks with $N = 30$.

260 Another important feature of collaboration networks is that they are not
 261 static. Scientists sometimes change their behavior, choosing to collaborate
 262 when they did not before and vice-versa. There are of course many possible
 263 ways to represent this. Following O’Connor (2017), Rubin and O’Connor
 264 (2018), and O’Connor et al. (2019), we suppose that scientists update their
 265 behavior using a rule known as “myopic best response”. This means that,
 266 in the first round of interaction, scientists choose a behavior at random. So
 267 a third of scientists plays *Low*, a third plays *Med*, and a third plays *High*.
 268 In each round thereafter, there is a small probability that a scientist updates
 269 their behavior. When a scientist updates their behavior, the scientist chooses
 270 the strategy that would have been a best response to the set of strategies
 271 that they encountered in the previous round. Scientists therefore update
 272 their behavior by best responding to previous plays but only keep a record
 273 of the most recent interactions.

274 Given our interest in the emergence of inequality in collaboration net-

275 works, we track how unequal the payoff distribution is. To do so, we use the
 276 Gini Index (GI). The GI measures the spread in a distribution. Although
 277 not entirely free of problems (Langel and Tillé, 2013), the GI is often used in
 278 economics to measure income and wealth inequality. It has also been applied
 279 to a variety of other contexts, such as in the study of biodiversity and enzyme
 280 selectivity (Wittebolle et al., 2009; Graczyk, 2007). The GI is given by:

$$GI = \frac{\sum_{i=1}^N \sum_{j=1}^N |\pi_i - \pi_j|}{2N \sum_{j=1}^N \pi_j} \quad (3)$$

281 where π_i and π_j are the payoffs that scientists i and j get from their collab-
 282 orations. The numerator is the mean absolute difference of the payoff dis-
 283 tribution and the denominator is twice the mean of the distribution. Since
 284 payoffs are always non-negative, the GI ranges from 0 (minimum) to 1 (max-
 285 imum) depending on the spread of the distribution. The GI thus measures
 286 the spread in the payoff distribution.

287 But we show below that it is possible for different aspects of inequality
 288 to come apart. For example, heterogeneity in the distribution of strategies
 289 can be low while payoff inequality is high (and vice versa). For this reason,
 290 we introduce another measure to track heterogeneity in the distribution of
 291 strategies: the Strategy Heterogeneity Index (SI). Since agents get the same
 292 payoff when both play *Med*, we define the SI as the overall frequency of
 293 agents who play any of the two extreme strategies (i.e., *Low* and *High*).
 294 The SI is therefore given by:

$$SI = f_L + f_H \quad (4)$$

295 where f_L and f_H give the frequency of agents who play *Low* and *High*,
 296 respectively. The SI ranges from 0 (minimum) to 1 (maximum), with 0
 297 indicating that everyone plays *Med* and 1 that no one plays *Med*. Unlike
 298 the GI , the SI therefore does not track the spread in the payoff distribution;
 299 it is instead a simple measure of how far the population deviates from the
 300 state in which everyone plays *Med*.

301 Having defined the structure of the collaboration network, the strategies
 302 that scientists in the collaboration network can adopt, the rule they use to up-
 303 date strategies, their payoffs, as well as two measures of inequality, we report
 304 our results in the next section. Pseudo-code, code for simulations, data, and
 305 scripts for analyses and figures are available anonymously at: <https://osf>.

306 [io/h6j75/?view_only=479ac3174b8c4f8e8b6e2de1af3e5abe](https://github.com/h6j75?view_only=479ac3174b8c4f8e8b6e2de1af3e5abe). Pseudo-code
307 is also available in the Appendix.

308 4 Results

309 Computer simulations show that collaboration networks reach an equilibrium
310 state in regular and scale-free networks. But regular and scale-free networks
311 arrive at different equilibria. In regular networks with $d = 2$ and $d = 5$,
312 the entire population comes to play *Med* when $L = 0.1$ (Figure 2, *left*). In
313 scale-free networks, however, only about 70% of the population plays *Med*
314 at equilibrium. Equilibria also differ when $L = 0.4$ (Figure 2, *right*). While
315 the entire population continues to play *Med* in regular networks with $d = 5$,
316 about 40% of the population comes to play *Med* in regular networks with
317 $d = 2$. In scale-free networks, the share of the population playing *Med*
318 is even smaller: about a third plays *Med*. The share of the population that
319 plays *Med* at equilibrium therefore depends on not only network topology,
320 but also average degree and value of L . (Since $L = 1 - H$, it does not matter
321 whether we track L or H ; we focus on L when presenting results.)

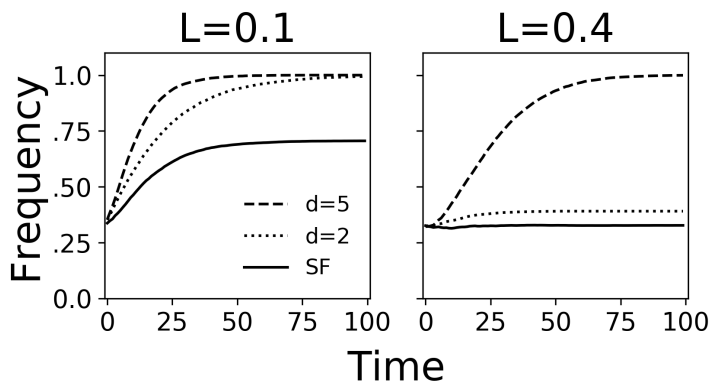


Figure 2: **Frequency of *Med* over time.** *Left*: when $L = 0.1$, *Med* takes over regular networks with $d = 2$ (*dotted*) and $d = 5$ (*dashed*); the equilibrium frequency of *Med* is 0.7 in scale-free networks (*solid*). *Right*: when $L = 0.4$, *Med* takes over regular networks with $d = 5$ but the frequency of *Med* is 0.4 in regular networks with $d = 2$ and 0.33 in scale-free networks. Results are average of 100 runs, update probability equal to 0.1, and $N = 100$.

322 We also find that the equilibrium composition of scale-free networks varies

323 across values of L (Figure 3, *left*). When $L = 0.1$, 72% of the population
 324 play *Med*, while 19% play *Low* and 9% play *High*. With increasing values of
 325 L , the equilibrium frequency of *Med* goes down while the frequencies of *Low*
 326 and *High* go up. When $L = 0.4$, the frequency of *High* is higher than the
 327 frequency of *Low*: 40% of the population play *High*, while 35% play *Med*
 328 and 25% play *Low*. Depending on L , the population thus goes from having
 329 more agents who play *Low* than *High* to having more *High* than *Low*.

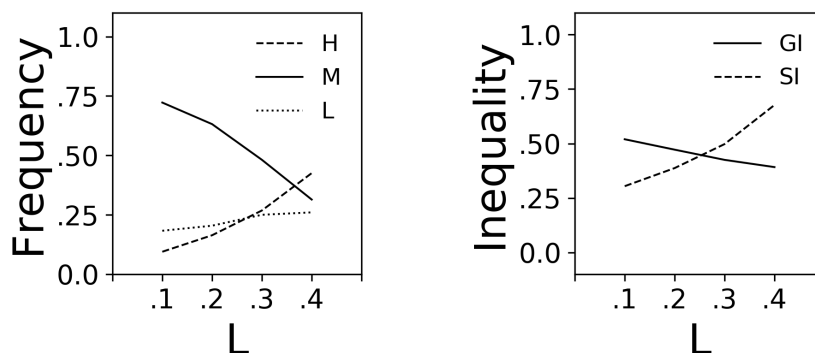


Figure 3: **Equilibrium Composition & Inequality.** *Left*: the equilibrium composition depends on L . *Right*: the Gini Index (GI) decreases with L , while the Strategy Heterogeneity Index (SI) increases with L . Results are average of 100 runs with 100 time steps, update probability equal to 0.1, and $N = 100$.

330 Next, we find that the payoff distribution becomes less unequal as L goes
 331 up (Figure 3, *right*). When $L = 0.1$, GI is about 0.52; when $L = 0.4$, GI
 332 is about 0.4. This is not very surprising given that higher (lower) values of
 333 L represent more egalitarian (elitist) communities. But the value of L has
 334 a very different effect on strategy heterogeneity: SI increases with L , with
 335 SI going from 0.3 when $L = 0.1$ to 0.66 when $L = 0.4$. These two measures
 336 also differ in that SI is more sensitive than GI to changes in the value of L :
 337 SI goes up by 120%, whereas GI goes down by 23%. As L increases, the
 338 population thus becomes less unequal with respect to payoff at the same time
 339 that it becomes a lot more heterogeneous with respect to its composition. In
 340 other words, payoff inequality and strategy heterogeneity come apart.

341 To better understand what factor(s) could be driving and maintaining
 342 payoff inequality and strategy heterogeneity, we consider how an agent's

343 strategy depends on the position that they occupy in the collaboration net-
 344 work. In particular, we compare the degree of agents who play *Low* with
 345 those who play *High* (Figure 4, *left*). When $L = 0.1$, agents playing *High*
 346 tend to have a higher average degree than agents playing *Low*: the former
 347 have about 3.6 collaborators on average, while the latter have about 1.24.
 348 But when $L = 0.4$, the pattern is reversed: agents playing *Low* tend to have
 349 about 3 collaborators, while agents playing *High* have around 1.36. When L
 350 is low, those who play *High* therefore tend to be well-connected agents; when
 351 L is high, it is those playing *Low* who are more likely to be well-connected.
 352 Inspection of a representative network at equilibrium illustrates this point
 353 (Figure 4, *right*). When $L = 0.1$, agents playing *Low* tend to occupy more
 354 peripheral nodes than agents playing *High*. Given that agents are identical
 355 except for the position that they occupy in the collaboration network, this
 356 suggests that it is the structure of the network that drives and maintains
 357 inequality in our model.

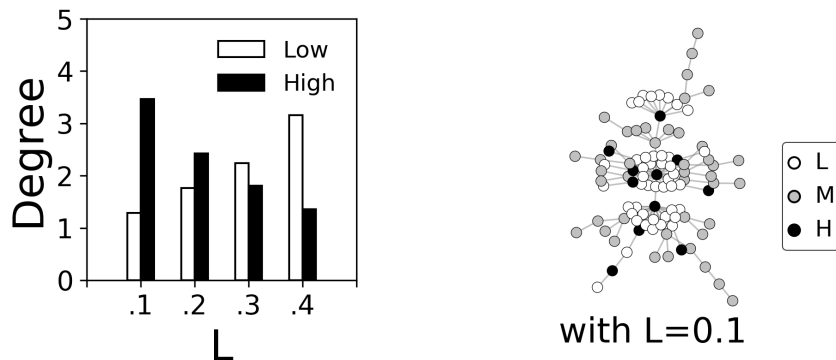


Figure 4: **Degree inequality in model networks.** *Left*: When L is low, the average degree of those playing *High* is higher than the average degree of those playing *Low*; the pattern is reversed when L is high. Results are average of 100 runs with 100 times steps, update probability equal to 0.1, and $N = 100$. *Right*: Population composition after 100 rounds of interactions in a scale-free collaboration network with $L = 0.1$.

358 But the structure of the collaboration network in our model is simply
 359 due to the preferential-attachment model. Although this model of network
 360 formation gives rise to a degree distribution that is known to resemble the
 361 degree distribution of real-world collaboration networks, it is clearly an ide-

362 alization. For one, scientists do not always choose who to collaborate with
 363 on the basis of how many collaborations potential coworkers already have—
 364 among myriad other factors, geographical proximity, institutional affiliation,
 365 and personality quirks can also play a role. To examine whether the in-
 366 equality we observe in our model might arise in the real world, we study the
 367 same dynamics of collaboration on two well-known and publicly available
 368 collaboration networks: the *GR-QC* and the *Erdos* collaboration network.
 369 The *GR-QC* collaboration network includes the authors of papers on general
 370 relativity and quantum cosmology posted to the pre-print repository arXiv
 371 between 1993 and 2003 (Leskovec et al., 2007). The *Erdos* collaboration net-
 372 work covers all papers written by the extremely prolific mathematician Paul
 373 Erdős, his co-authors, and their co-authors (Batagelj and Mrvar, 2000).

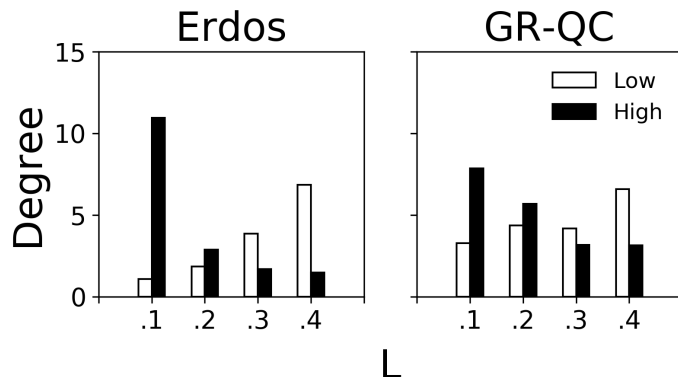


Figure 5: **Degree inequality in real-world networks.** In the *Erdos* ($N = 4, 158$; *left*) and the *GR-QC* ($N = 5, 094$; *right*) collaboration networks, the average degree of agents who play *High* is higher than the average degree of agents who play *Low* when L is low; the pattern is reversed when L is high. Results are average of 100 runs with 100 time steps and update probability equal to 0.1.

374 We obtain similar results from simulations of a population of agents play-
 375 ing the mini-Nash demand game with myopic best response on the *GR-QC*
 376 and the *Erdos* collaboration networks (Figure 5). In particular, the average
 377 degree is higher for agents playing *High* than for agents playing *Low* when
 378 L is low but the pattern is reversed when L is high. When $L = 0.1$, scientists
 379 in *GR-QC* who play *Low* have about 3.1 collaborators on average, while sci-
 380 entists who play *High* have about 7.9 collaborators. A similar pattern holds

381 in *Erdos*: when $L = 0.1$, scientists playing *Low* have a single collaborator on
 382 average but scientists playing *High* have about 10.9 collaborators. As L goes
 383 up, this difference decreases at first and eventually reverses. When $L = 0.4$,
 384 scientists in *GR-QC* who play *Low* have about 6.37 collaborators on average,
 385 while scientists playing *High* have about 2.94. Similarly, scientists in
 386 *Erdos* who play *Low* have 7 collaborators on average, while scientists playing
 387 *High* have about 1.46. Network structure therefore drives the emergence of
 388 inequality in both networks, although the effect is especially pronounced in
 389 *Erdos*.

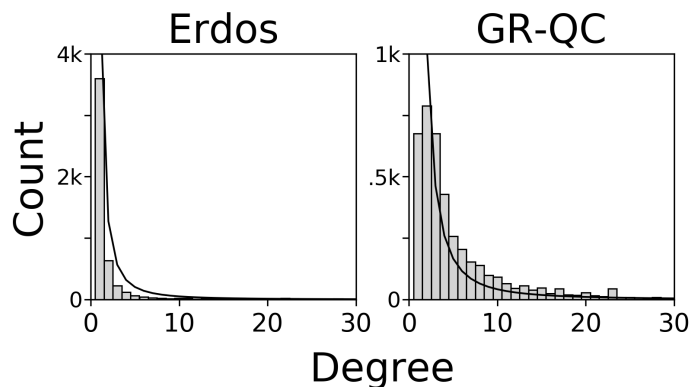


Figure 6: **Degree distribution in model and two real-world networks.**
Left: the degree distribution given by $P(d) = N \cdot d^{-\gamma}$ with $\gamma = 2$ (*solid line*) approximates the degree distribution in the *Erdos* collaboration network ($N = 4,158$). *Right:* the same expression approximates the observed degree distribution in the *GR-QC* collaboration network ($N = 5,094$). Grey bars show empirical degree distribution.

390 It is also worth reiterating that the degree distribution of scale-free net-
 391 works where inequality arises is similar to that of real-world collaboration
 392 networks. As already noted, the degree distribution of indefinitely large
 393 scale-free networks is given by $P(d) \sim d^{-\gamma}$. Empirical studies find that val-
 394 ues of γ for real-world collaboration networks often range between values of
 395 1 and 3, depending on dataset and scientific discipline (Barabási et al., 2002;
 396 Albert and Barabási, 2002). Indeed, this expression approximates quite well
 397 the degree distribution of both the *Erdos* and the *GR-QC* collaboration net-
 398 works (Figure 6). Considering that the preferential-attachment model was
 399 built to fit the scale-free degree distribution of real-world networks, this is not

400 very surprising. But it serves as a reminder that the inequality we observe
401 in our model is the product of a realistic network structure.

402 5 Discussion

403 Our model shows that the structure of collaboration networks can give rise to
404 inequality even in the absence of social categories. In particular, our model
405 shows that inequality in the payoff distribution and heterogeneity in the
406 strategy profile of the population arises and persists in collaboration networks
407 with a heterogeneous degree distribution. Our model also shows that this
408 is so across the full range of values for L —a parameter that controls how
409 elitist or egalitarian the scientific community tends to be. Furthermore, our
410 model highlights that inequality is not a one-dimensional concept: different
411 values of L affect different measures of inequality differently, with inequality
412 in the payoff distribution (GI) being high when heterogeneity in the strategy
413 profile (SI) is low and vice-versa.

414 These results stand in contrast to previous models showing that popula-
415 tion structure can promote an even allocation of resources in the mini-Nash
416 demand game. For example, Alexander and Skyrms (1999) and Alexander
417 (2000) show that spatial structure makes it very likely that a population will
418 converge on the fair equilibrium. But this is due to the fact that spatial
419 organization is a form of population structure where every agents interacts
420 with four neighbors and there is no variation in the degree distribution. When
421 population structure leads many to interact with few and few to interact with
422 many, our model shows that the resulting heterogeneous degree distribution
423 can promote unequal outcomes.

424 Our model thus adds to a growing body of work showing that a hetero-
425 geneous degree distribution can give rise to inequalities in strategic settings.
426 In a network model of the Prisoner’s Dilemma, for example, Du et al. (2008),
427 find that a heterogeneous degree distribution favors the spread of coopera-
428 tion but that it also promotes an unequal payoff distribution. In public goods
429 games, network heterogeneity induces diversity in group size and thus pro-
430 motes contributions to the public good (Santos et al., 2006, 2008). But net-
431 work heterogeneity can also lead to unequal outcomes in public good games,
432 as the proliferation of altruistic behaviors ends up harming some individuals
433 (McAvoy et al., 2020).

434 Our model also reveals two “regimes” in the emergence of inequality in

435 collaboration networks. One regime is when L is low. In this case, poorly
436 connected scientists in the periphery of the collaboration network play *Low*,
437 while their well-connected collaborators play *High*. The other regime is when
438 L is high. In this case, well-connected scientists play *Low*, while their poorly
439 connected collaborators play *High*. An analogous pattern is apparent in the
440 way that the Red King/Queen effect leads to inequality in the mini-Nash
441 bargaining game with co-evolving groups of different sizes (Bruner, 2019;
442 O’Connor, 2019; O’Connor, 2017). When L is high, the Red King effect
443 leads the minority to get less than the majority. When L is low, the Red
444 Queen effect kicks in and the minority gets more than the majority.

445 Despite this superficial similarity, the mechanism driving the emergence
446 of inequality in our model is not the same as in the Red King/Queen. First,
447 the Red King/Queen depends on the minority adapting more quickly to
448 the strategy of the majority. In contrast, the update rule we use is the
449 myopic best response. Strictly speaking, the myopic best response is not
450 an evolutionary update rule because agents do not update their behavior
451 by copying the behavior of others. So it is not a difference in evolutionary
452 tempo that drives inequality in our model. Second, the Red King/Queen
453 relies on there being two groups, groups having different sizes, and individuals
454 conditionalizing their behavior on the group membership of others. In our
455 model, however, the mechanism that gives rise to inequality does not depend
456 on a categorical distinction between groups. In fact, there is no partition of
457 the population into groups at all—let alone groups of different sizes. Third,
458 the Red King/Queen effect causes the minority groups to be at a disadvantage
459 when L is high and thus when payoff inequality is low. But in our model
460 those who are poorly connected end up at a disadvantage when L is low and
461 payoff inequality is high. For all these reasons, the mechanism leading to
462 inequality in our model is not the same as the Red King/Queen.

463 So what explains the two regimes of inequality that we observe in our
464 model? Since the update rule we use is the myopic best response, to answer
465 this question we follow Rubin and O’Connor’s (2018, pp. 386-8) account of
466 how discrimination arises in their model and consider the probability that
467 a strategy is a best response.⁴ A strategy is a best response if there is no
468 other strategy that would yield a higher payoff given the strategies that other
469 agents play in the previous round. The probability that a particular strategy
470 is a best response thus depends on the probability with which other agents

⁴We thank an anonymous referee for raising this point.

471 choose each strategy. For an agent who only interacts with one other agent,
 472 the probability that the strategy *Low*, *Med*, or *High* is a best response
 473 is just the probability with which the agent encounters another agent who
 474 plays *High*, *Med*, or *Low*. Initially, agents choose a strategy at random.
 475 The initial probability that each strategy is a best response is thus $\frac{1}{3}$.

476 In scale-free networks, some agents do interact with only one other agent.
 477 But other agents interact with many more. In such cases, the probability that
 478 a strategy is a best response can be found in three steps. The first step is to
 479 determine what strategy is a best response to every possible combination of
 480 strategies that other agents may choose. The second step is to calculate the
 481 probability with which each one of these combinations of strategies occurs.
 482 The third step is to compute the probability that a strategy is a best response
 483 by summing over the probabilities of every combination of strategies to which
 484 the strategy in question is a best response. Assuming that agents pick a
 485 strategy at random, as they do at first, the probability that *Low*, *Med*, or
 486 *High* is a best response is shown in Figure 7.

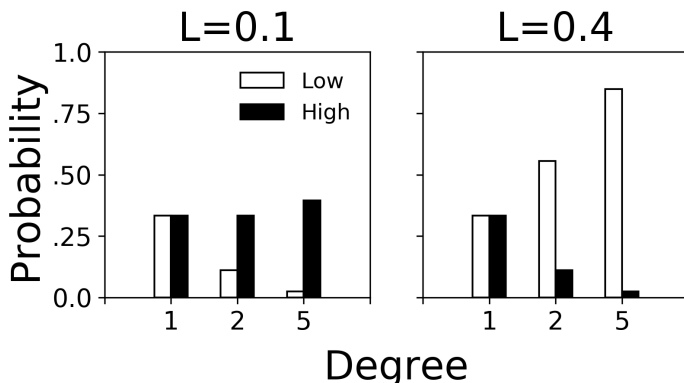


Figure 7: **Initial probabilities that *Low* and *High* is a best response.**
Left: initial probability that *Low* and *High* are a best response for $d = 1$,
 $d = 2$, and $d = 5$ when $L = 0.1$. *Right:* initial probability that *Low* and
High are a best response for $d = 1$, $d = 2$, and $d = 5$ when $L = 0.4$.

487 Notice that the probability that a strategy is a best response depends on
 488 degree. As already noted, each strategy is a best response with probability $\frac{1}{3}$
 489 when an agent interacts with only one other agent—and this is so regardless
 490 of L . But when an agent interacts with more than one agent, the probability
 491 that a strategy is a best response depends on how many other agents they

492 interact with. When $L = 0.1$, for example, the probability that *Low* is a best
493 response for an agent who interacts with two other agents is about 0.11. But
494 the probability that *Low* is a best response for an agent who interacts with
495 five other agents is only 0.025. When $L = 0.4$, the probability that *Low* is a
496 best response for an agent who interacts with two other agents is about 0.55.
497 But the probability that *Low* is a best response for an agent who interacts
498 with five other agents is about 0.85.

499 This allows us to gain some insight into the two regimes for the emergence
500 of inequality in our model. Consider two groups of agents: poorly connected
501 agents with $d = 1$, and well-connected agents with $d \geq 5$. When $L = 0.1$, the
502 initial probability that *Low* or *High* is a best response for poorly connected
503 agents is one third. But for well-connected agents the initial probability that
504 *High* is a best response is a lot higher than the initial probability that *Low*
505 is a best response. This is because the relative payoff to *High* is relatively
506 high, so well-connected individuals respond best by “sticking to their guns”
507 and making a *High* demand that yields a large increase in payoff. For this
508 reason, well-connected agents tend to play *High* and end up at an advantage
509 when L is low; at the same time, poorly connected agents tend to play *Low*
510 and end up at a disadvantage. When $L = 0.4$, the initial probability that
511 *Low* or *High* is a best response for poorly connected agents is again one
512 third. For well-connected agents, however, the initial probability that *Low* is
513 a best response is now a lot higher than the initial probability that *High* is
514 a best response. This is because the relative payoff to *Low* is relatively high,
515 so well-connected individuals respond best by playing it safe and making a
516 *Low* demand instead of holding out for what would be a small increase in
517 payoff. Well-connected individuals therefore tend to play *Low* and end up
518 at a disadvantage when L is high, while poorly connected agents play *High*
519 and end up at an advantage. The two regimes of inequality we observe in
520 scale-free networks is thus due to differences in the initial probability that a
521 strategy is a best response.⁵

522 From a social-epistemological perspective, this raises a series of impor-
523 tant questions about the structure of collaboration networks. Well-connected

⁵The initial probabilities that either *Low* or *High* is a best response is higher when $L = 0.4$ than when $L = 0.1$ for $d \geq 2$. This helps explain why a smaller share of the population comes to play *Med* in scale-free networks and regular networks with $d = 2$ when L is high. In regular networks with $d = 5$, the initial probability that *Low* is a best response is so high that the population quickly becomes saturated with *Low*. This decreases the probability that *Low* is a best response and allows *Med* to take over.

524 scientists are more likely to play *Low* and end up at a disadvantage when
525 L is high. This means that well-connected scientists are at a disadvantage
526 in egalitarian communities where payoff inequality is low. Poorly connected
527 scientists, however, are more likely to play *Low* and thus end up at a disad-
528 vantage when L is low. Low values of L correspond to elitist communities
529 where payoff inequality is high. Our model therefore raises the specter of a
530 two-fold harm: low values of L put poorly connected scientists at a disad-
531 vantage when doing so is particularly harmful.

532 The two-fold harm of structural inequality is all the more worrisome
533 because members of minority or underrepresented groups are often poorly
534 connected in real-world collaboration networks. Female scientists, for exam-
535 ple, have fewer collaborators than their male colleagues (Araujo et al., 2017;
536 Abramo et al., 2009). Black scientists also have fewer collaborators, at least
537 in some disciplines (Del Carmen and Bing, 2000). When payoff inequality
538 is especially high, the two-fold harm is likely to arise and members of these
539 groups might therefore be at a disadvantage. To make matters worse, im-
540 plicit and explicit biases linked to social categories might only exacerbate
541 the problem: prejudice and discrimination tends to put those groups at a
542 disadvantage who are already vulnerable due to the position that they oc-
543 cupy in the collaboration network. For example, if scientists choose what
544 collaborations to enter on the basis of biases against visible group markers,
545 then biases and social categories might contribute to the formation of collab-
546 oration networks where pernicious forms of structural inequality are likely to
547 emerge.

548 6 Conclusion

549 Philosophers have long worried that implicit and explicit biases are inevitable
550 in science and that they often contribute to various forms of epistemic in-
551 justice (Longino, 1990; Fricker, 2007). In recent years, formal models in
552 philosophy of science have further shown that it is possible for discrimina-
553 tory norms to lead to an unequal allocation of epistemic credit even when
554 there are no biases (O'Connor and Bruner, 2019; Rubin and O'Connor, 2018;
555 O'Connor et al., 2019). But models proposed so far account for these wor-
556 risome patterns in research by positing the existence of social categories.
557 Although biases and social categories remain a source of concern, we show
558 that unequal outcomes are possible even in the absence of social categories:

559 when scientists bargain with collaborators in a scale-free network, inequality
560 arises simply because of the structure of the collaboration network. We also
561 bring empirical considerations to bear on models of the social organization of
562 science by showing that structural inequality can likewise arise in real-world
563 collaboration networks (cf. Martini and Pinto, 2017).

564 It is important to keep in mind, however, that our model makes several
565 simplifying assumptions. First, we assume that scientist play the same strat-
566 egy with all their collaborations. This is unlikely to hold in reality since
567 scientists often negotiate different arrangements with different collaborators.
568 Second, we consider a dynamic population of scientists who change their
569 strategies over time but assume that the structure of the collaboration net-
570 work is static. This is not the case in the real world where scientists can
571 not only update their behavior, but also adjust their social ties. Third, we
572 assume that all scientists are equally competent. This is again unrealistic be-
573 cause scientists often differ with respect to how productive they are. Fourth,
574 we assume that scientists update their strategy by myopic best response.
575 This is a reasonable assumption but update rules based on imitation are
576 also plausible. While these simplifying assumptions allow us to isolate and
577 better understand an important phenomenon, it would be interesting to re-
578 lax these assumptions. Future work could therefore consider collaboration
579 networks where scientists pursue different strategies with different collabo-
580 rators, change who they interact with over time, differ with respect to how
581 productive they are, or update their strategy according to different rules.

582 **Appendix**

583 We use a simple program to simulate the behavior of agents in a network
584 who interact with their neighbors by playing the mini-Nash demand game.
585 In pseudo-code, the program proceeds as follows:

```
586 _____  
587  
588 FOR each Network Topology, DO:  
589     FOR each Agent, DO:  
590         Choose Demand at random from options L, M, and H  
591         FOR each Time Step, DO:  
592             FOR each Agent, DO:  
593                 Get Agent's Demand  
594                 Get Demand for each of Agent's neighbors  
595                 Get Agent's payoff based on own Demand and neighbors' Demands  
596             With probability 0.1, DO:  
597                 Find Agent's Best Response in previous Time Step  
598                 Update Agent's Demand  
599 _____  
600
```

601 **References**

- 602 Abelson, J. S., N. Z. Wong, M. Symer, G. Eckenrode, A. Watkins, and H. L.
603 Yeo (2018). Racial and ethnic disparities in promotion and retention of
604 academic surgeons. *The American Journal of Surgery* 216(4), 678–682.
- 605 Abramo, G., C. D’Angelo, and A. Caprasecca (2009). The contribution of
606 star scientists to overall sex differences in research productivity. *Sciento-*
607 *metrics* 81(1), 137–156.
- 608 Albert, R. and A.-L. Barabási (2002). Statistical mechanics of complex net-
609 works. *Reviews of Modern Physics* 74(1), 47.
- 610 Alexander, J. M. (2000). Evolutionary explanations of distributive justice.
611 *Philosophy of Science* 67(3), 490–516.
- 612 Alexander, J. M. and B. Skyrms (1999). Bargaining with neighbors: Is justice
613 contagious? *The Journal of Philosophy* 96(11), 588–598.
- 614 Allison, P. D., J. S. Long, and T. K. Krauze (1982). Cumulative advantage
615 and inequality in science. *American Sociological Review* 47(5), 615–625.
- 616 Allison, P. D. and J. A. Stewart (1974). Productivity differences among
617 scientists: Evidence for accumulative advantage. *American sociological*
618 *review* 39(4), 596–606.
- 619 Araujo, E. B., N. A. Araújo, A. A. Moreira, H. J. Herrmann, and J. S.
620 Andrade Jr (2017). Gender differences in scientific collaborations: Women
621 are more egalitarian than men. *PloS one* 12(5), e0176791.
- 622 Axtell, R. L., J. M. Epstein, and H. P. Young (2001). The emergence of
623 classes in a multiagent bargaining model. *Social dynamics* 27, 191–211.
- 624 Barabási, A.-L. and R. Albert (1999). Emergence of scaling in random net-
625 works. *Science* 286(5439), 509–512.
- 626 Barabási, A.-L., H. Jeong, Z. Néda, E. Ravasz, A. Schubert, and T. Vicsek
627 (2002). Evolution of the social network of scientific collaborations. *Physica*
628 *A: Statistical mechanics and its applications* 311(3-4), 590–614.
- 629 Barabási, A.-L. and Z. N. Oltvai (2004). Network biology: understanding
630 the cell’s functional organization. *Nature Reviews Genetics* 5(2), 101–113.

- 631 Batagelj, V. and A. Mrvar (2000). Some analyses of erdos collaboration
632 graph. *Social Networks* 22(2), 173–186.
- 633 Beaver, D. d. (2004). Does collaborative research have greater epistemic
634 authority? *Scientometrics* 60(3), 399–408.
- 635 Bergstrom, C. T. and M. Lachmann (2003). The red king effect: when the
636 slowest runner wins the coevolutionary race. *Proceedings of the National*
637 *Academy of Sciences* 100(2), 593–598.
- 638 Binmore, K. G. (1998). *Game theory and the social contract: just playing*,
639 Volume 2. MIT press.
- 640 Bol, T., M. de Vaan, and A. van de Rijt (2018). The matthew effect in science
641 funding. *Proceedings of the National Academy of Sciences* 115(19), 4887–
642 4890.
- 643 Bruner, J. P. (2019). Minority (dis) advantage in population games. *Syn-*
644 *these* 196(1), 413–427.
- 645 Clauset, A., S. Arbesman, and D. B. Larremore (2015). Systematic inequality
646 and hierarchy in faculty hiring networks. *Science advances* 1(1), e1400005.
- 647 Del Carmen, A. and R. L. Bing (2000). Academic productivity of african
648 americans in criminology and criminal justice. *Journal of Criminal Justice*
649 *Education* 11(2), 237–249.
- 650 Du, W.-B., H.-R. Zheng, and M.-B. Hu (2008). Evolutionary prisoner’s
651 dilemma game on weighted scale-free networks. *Physica A: Statistical Me-*
652 *chanics and its Applications* 387(14), 3796–3800.
- 653 Feldon, D. F., J. Peugh, M. A. Maher, J. Roksa, and C. Tofel-Grehl (2017).
654 Time-to-credit gender inequities of first-year phd students in the biological
655 sciences. *CBE—Life Sciences Education* 16(1), ar4.
- 656 Fricker, M. (2007). *Epistemic injustice: Power and the ethics of knowing*.
657 Oxford University Press.
- 658 Gabbidon, S. L., H. T. Greene, and K. Wilder (2004). Still excluded? an
659 update on the status of african american scholars in the discipline of crim-
660 inology and criminal justice. *Journal of Research in Crime and Delin-*
661 *quency* 41(4), 384–406.

- 662 Graczyk, P. P. (2007). Gini coefficient: a new way to express selectivity of
663 kinase inhibitors against a family of kinases. *Journal of Medicinal Chem-*
664 *istry* 50(23), 5773–5779.
- 665 Han, S.-K. (2003). Tribal regimes in academia: A comparative analysis of
666 market structure across disciplines. *Social networks* 25(3), 251–280.
- 667 Henriksen, D. (2016). The rise in co-authorship in the social sciences (1980–
668 2013). *Scientometrics* 107(2), 455–476.
- 669 Holman, B. and J. Bruner (2017). Experimentation by industrial selection.
670 *Philosophy of Science* 84(5), 1008–1019.
- 671 Hopkins, A. L., J. W. Jawitz, C. McCarty, A. Goldman, and N. B. Basu
672 (2013). Disparities in publication patterns by gender, race and ethnicity
673 based on a survey of a random sample of authors. *Scientometrics* 96(2),
674 515–534.
- 675 Kitcher, P. (1990). The division of cognitive labor. *The Journal of Philoso-*
676 *phy* 87(1), 5–22.
- 677 Kummerfeld, E. and K. J. Zollman (2015). Conservatism and the scientific
678 state of nature. *The British Journal for the Philosophy of Science* 67(4),
679 1057–1076.
- 680 Langel, M. and Y. Tillé (2013). Variance estimation of the gini index: re-
681 visiting a result several times published. *Journal of the Royal Statistical*
682 *Society: Series A (Statistics in Society)* 176(2), 521–540.
- 683 Larivière, V., C. Ni, Y. Gingras, B. Cronin, and C. R. Sugimoto (2013). Bib-
684 liometrics: Global gender disparities in science. *Nature News* 504(7479),
685 211.
- 686 Lee, S. and B. Bozeman (2005). The impact of research collaboration on
687 scientific productivity. *Social Studies of Science* 35(5), 673–702.
- 688 Leskovec, J., J. Kleinberg, and C. Faloutsos (2007). Graph evolution: Den-
689 sification and shrinking diameters. *ACM transactions on Knowledge Dis-*
690 *covery from Data (TKDD)* 1(1), 2–es.
- 691 Longino, H. E. (1990). *Science as social knowledge: Values and objectivity*
692 *in scientific inquiry*. Princeton University Press.

- 693 Lusseau, D. (2003). The emergent properties of a dolphin social network.
694 *Proceedings of the Royal Society of London. Series B: Biological Sci-*
695 *ences* 270(suppl_2), S186–S188.
- 696 Martini, C. and M. F. Pinto (2017). Modeling the social organization of
697 science. *European Journal for Philosophy of Science* 7(2), 221–238.
- 698 McAvoy, A., B. Allen, and M. A. Nowak (2020). Social goods dilemmas in
699 heterogeneous societies. *Nature Human Behaviour* 4(8), 819–831.
- 700 Melin, G. and O. Persson (1996). Studying research collaboration using co-
701 authorships. *Scientometrics* 36(3), 363–377.
- 702 Merton, R. K. (1968). The matthew effect in science: The reward and com-
703 munication systems of science are considered. *Science* 159(3810), 56–63.
- 704 Mohseni, A., C. O’Connor, and H. Rubin (2019). On the emergence of
705 minority disadvantage: testing the cultural red king hypothesis. *Synthese*,
706 1–23.
- 707 Nash, J. F. J. (1950). The bargaining problem. *Econometrica* 18(2), 155–162.
- 708 Newman, M. E. (2001). The structure of scientific collaboration networks.
709 *Proceedings of the National Academy of Sciences* 98(2), 404–409.
- 710 Newman, M. E. (2004). Coauthorship networks and patterns of scientific
711 collaboration. *Proceedings of the National Academy of Sciences* 101(suppl
712 1), 5200–5205.
- 713 Nielsen, M. W. and J. P. Andersen (2021). Global citation inequality is on
714 the rise. *Proceedings of the National Academy of Sciences* 118(7), 1–10.
- 715 O’Connor, C. (2019). *The origins of unfairness: Social categories and cultural*
716 *evolution*. Oxford University Press, USA.
- 717 O’Connor, C. (2017). The cultural red king effect. *The Journal of Mathe-*
718 *matical Sociology* 41(3), 155–171.
- 719 O’Connor, C., L. K. Bright, and J. P. Bruner (2019). The emergence of
720 intersectional disadvantage. *Social Epistemology* 33(1), 23–41.

- 721 O'Connor, C. and J. Bruner (2019). Dynamics and diversity in epistemic
722 communities. *Erkenntnis* 84(1), 101–119.
- 723 Petersen, A. M., S. Fortunato, R. K. Pan, K. Kaski, O. Penner, A. Rungi,
724 M. Riccaboni, H. E. Stanley, and F. Pammolli (2014). Reputation and
725 impact in academic careers. *Proceedings of the National Academy of Sci-*
726 *ences* 111(43), 15316–15321.
- 727 Rosenstock, S., J. Bruner, and C. O'Connor (2017). In epistemic networks,
728 is less really more? *Philosophy of Science* 84(2), 234–252.
- 729 Rubin, H. and C. O'Connor (2018). Discrimination and collaboration in
730 science. *Philosophy of Science* 85(3), 380–402.
- 731 Santos, F. C., J. M. Pacheco, and T. Lenaerts (2006). Evolutionary dynamics
732 of social dilemmas in structured heterogeneous populations. *Proceedings*
733 *of the National Academy of Sciences* 103(9), 3490–3494.
- 734 Santos, F. C., M. D. Santos, and J. M. Pacheco (2008). Social diver-
735 sity promotes the emergence of cooperation in public goods games. *Na-*
736 *ture* 454(7201), 213–216.
- 737 Skyrms, B. (1996). *Evolution of the social contract*. Cambridge University
738 Press.
- 739 Skyrms, B. and K. J. Zollman (2010). Evolutionary considerations in the
740 framing of social norms. *Politics, philosophy & economics* 9(3), 265–273.
- 741 Weatherall, J. O., C. O'Connor, and J. P. Bruner (2020). How to beat science
742 and influence people: policymakers and propaganda in epistemic networks.
743 *The British Journal for the Philosophy of Science* 71(4), 1157–1186.
- 744 Weisberg, M. and R. Muldoon (2009). Epistemic landscapes and the division
745 of cognitive labor. *Philosophy of Science* 76(2), 225–252.
- 746 West, J. D., J. Jacquet, M. M. King, S. J. Correll, and C. T. Bergstrom
747 (2013). The role of gender in scholarly authorship. *PloS one* 8(7), e66212.
- 748 Wittebolle, L., M. Marzorati, L. Clement, A. Balloi, D. Daffonchio,
749 K. Heylen, P. De Vos, W. Verstraete, and N. Boon (2009). Initial
750 community evenness favours functionality under selective stress. *Na-*
751 *ture* 458(7238), 623–626.

- 752 Witteman, H. O., M. Hendricks, S. Straus, and C. Tannenbaum (2019). Are
753 gender gaps due to evaluations of the applicant or the science? a natural
754 experiment at a national funding agency. *The Lancet* 393(10171), 531–540.
- 755 Wray, K. B. (2002). The epistemic significance of collaborative research.
756 *Philosophy of Science* 69(1), 150–168.
- 757 Zollman, K. J. (2007). The communication structure of epistemic communi-
758 ties. *Philosophy of Science* 74(5), 574–587.
- 759 Zollman, K. J. (2010). The epistemic benefit of transient diversity. *Erkennt-*
760 *nis* 72(1), 17.